

Factors related to glucose-lowering efficacy of oral
antidiabetics: a systematic review and meta-analysis
focusing on ethnicity and study regions

Kayo Nagaki (Fujita)

Department of Clinical Medicine (Pharmaceutical Medicine)
Kitasato University Graduate School of Pharmaceutical Sciences

5-9-1 Shirokane, Minato-ku, Tokyo, 108-8641, Japan

Abstract

Type 2 diabetes is a group of metabolic diseases characterized by hyperglycemia resulting from lack of insulin action. The mechanism for a lack of insulin action has two components: impaired insulin secretion and decreased insulin sensitivity (insulin resistance) in organs on which insulin acts. Both are components of the pathogenesis of type 2 diabetes. In general, insulin resistance plays a key role causing type 2 diabetes in Caucasians, whereas impaired insulin secretion plays a key role in the development of type 2 diabetes in East Asians.

Several systematic reviews and meta-analyses have been conducted including an analysis to investigate the difference between ethnic groups in glucose lowering efficacy of dipeptidyl peptidase-4 (DPP-4) inhibitors. However, no consistent result has been found to date. Thus, in this first research, we performed a systematic review and meta-analysis of randomized controlled studies and conducted meta-regression analyses to assess the correlation between the glucose-lowering efficacy of DPP-4 inhibitors in patients with type 2 diabetes and potential factors including ethnicity of study subjects and study regions; in particular, Japanese were dealt separately from other Asians. The finding of the first research suggested that a higher baseline Hemoglobin A1c (HbA1c), studies in Japanese subjects, and studies conducted in Japan are factors of greater efficacy of DPP-4 inhibitors on weighted mean difference (WMD) in the change of HbA1c from baseline.

As for the sodium glucose co-transporter-2 (SGLT-2) inhibitor, systematic reviews and meta-analysis to investigate the difference in efficacy and safety of SGLT-2 inhibitors between different ethnic groups considering the pathophysiology has not been conducted so far. The reason might be its insulin independent mechanism of action.

Thus, as a secondary research, we performed a systematic review and meta-analysis of randomized controlled studies and conducted meta-regression analyses to assess the correlation between the glucose-lowering efficacy of SGLT-2 inhibitors in patients with type 2 diabetes and potential factors including ethnicity of study subjects and study regions. The finding of the second research suggested that higher fasting plasma glucose is a factor of greater efficacy of SGLT-2 inhibitors on WMD in the change of HbA1c from baseline.

Our finding suggests that paying attention to baseline HbA1c for DPP-4 inhibitors and baseline fasting plasma glucose for SGLT-2 inhibitors in planning and conducting clinical studies, and in comparing the data with other clinical studies is important. With regards to the ethnicity of study subjects, our finding indicates that differences in the contribution of the insulin secretory defect and the insulin resistance in the pathophysiology of type 2 diabetes between Asians and non-Asians do not affect the response to active oral hypoglycemic agents which have insulin independent mechanism.

In type 2 diabetes area, placebo effect has rarely been discussed. The third research focus, therefore, was to identify the factors that contribute to placebo effect in clinical studies for both DPP-4 inhibitor and SGLT-2 inhibitor by meta-regression analyses. The finding of the third research suggested that study conduct in Asia is a factor of larger HbA1c change from baseline in the placebo group. This differential HbA1c response in the placebo arm should be taken into consideration when comparing the data based on different clinical practice in each region. In the future, it is desirable to evaluate the influence of different study regions and/or subject ethnicity on efficacy and safety in multiregional studies.

Table of Contents

Abstract	i
Table of Contents	iii
List of Tables	iv
List of Figures	v
Abbreviations	vi
1. Introduction 1	1
2. Research 1 (DPP-4 inhibitor)	5
2.1 Objectives	5
2.2 Method	6
2.3 Result	9
2.4 Discussion	21
3. Research 2 (SGLT-2 inhibitor)	25
3.1 Objectives	25
3.2 Method	26
3.3 Result	30
3.4 Discussion	40
4. Research 3 (Placebo effect)	43
4.1 Objectives	43
4.2 Method	43
4.3 Result	44
4.4 Discussion	48
5. Overall Discussion	50
6. Conclusion	52
References	53
Acknowledgement	56
Appendix 1	57
Appendix 2	67

List of Tables

Table 1 Studies included in the analysis

Table 2 Three classifications for studies

Table 3 Results of multivariate meta-regression analysis

Table 4 Studies included in the analysis

Table 5 Three classifications for studies

Table 6 Results of multivariate meta-regression analysis

Table 7 Results of multivariate meta-regression analysis (*DPP-4 inhibitor study*)

Table 8 Results of multivariate meta-regression analysis (*SGLT-2 inhibitor study*)

List of Figures

Figure 1 Flow diagram of study selection

Figure 2 Funnel plot of weighted mean difference for HbA1c and standard error

Figure 3 Flow diagram of study selection

Figure 4 Funnel plot of weighted mean difference for HbA1c and standard error

Abbreviations

BMI	body mass index
DPP-4	dipeptidyl peptidase-4
GFR	glomerular filtration rate
HbA1c	Hemoglobin A1c
JDS	Japan Diabetes Society
NGSP	National Glycohemoglobin Standardization Program
PRISMA	Preferred Reporting Items for Systematic Reviews and Meta-Analyses
SGLT-2	sodium glucose co-transporter-2
WMD	weighted mean difference

1. Introduction

Type 2 diabetes is a group of metabolic diseases characterized by hyperglycemia resulting from lack of insulin action. The chronic hyperglycemia of diabetes is associated with long-term damage, dysfunction, and failure of various organs, especially the eyes, kidneys, nerves, heart, and blood vessels

The prevalence of diabetes has recently been increasing worldwide [1-3]. The number of adults with diabetes was estimated to be 415 million in 2015 and is expected to rise to 642 million by 2040 [3]. East Asian countries have also seen a continuous increase in diabetes [2,3]. According to a 1997 survey in Japan [4], total of “individuals strongly suspected of having diabetes” and “individuals in whom diabetes cannot be ruled out” were approximately 13.7 million, which increased to 20.5 million in 2012 [5].

The mechanism for a lack of insulin action has two components: impaired insulin secretion and decreased insulin sensitivity (insulin resistance) in organs on which insulin acts [6]. Impaired insulin secretion is a condition in which enough insulin is not supplied to satisfy the adequate action on the organs. Insulin resistance refers to a condition in which the normal function of insulin in peripheral tissues does not occur. Both are components of the pathogenesis of type 2 diabetes. In general, insulin resistance plays a key role causing type 2 diabetes in Caucasians [7], whereas impaired insulin secretion plays the key role in the development of type 2 diabetes in East Asians [8-11]. Besides the differences in pathogenesis between ethnic groups, there are also differences at the individual level. In the case of impaired insulin secretion, excessive lipid uptake occurs even in mild obesity, and when a lifestyle that is deficient in exercise is superimposed on this, insulin resistance, and thus type 2 diabetes, is more likely to

occur.

In controlling type 2 diabetes, lifestyle management focusing on diet and exercise is carried out first, and in case the target value for glycemic control is not achieved, drug treatment is started. Seven oral hypoglycemic agents are divided into three modes of actions: insulin sensitizing agents, insulin secretagogues, and carbohydrate absorption/ excretion-modulating agent. Biguanides and thiazolidines are classified as insulin sensitizing agents, sulfonylureas, fast-acting insulin secretagogues and dipeptidyl peptidase-4 (DPP-4) inhibitors are insulin secretagogues, and α -glucosidase inhibitor and sodium glucose co-transporter-2 (SGLT-2) inhibitors are carbohydrate absorption/excretion-modulating agents. Pharmacotherapy will be selected depending on the patient's etiology and patho-physiological stages (states) considering different modes of actions of each drug.

There is no apparent difference among the regions e.g., the United States, European countries, and Japan, regarding diagnosis and drug treatment options for type 2 diabetes. Hemoglobin A1c (HbA1c) is widely used internationally as an important index in the treatment of diabetes and a primary endpoint of efficacy in clinical studies. Based on the background such as the differences in pathogenesis and so on, however, it has been considered in Japan that extrapolation of multinational clinical study results by bridging strategy and/or conducting multinational studies is difficult in the development of an oral hypoglycemic agent for type 2 diabetes patients.

Several systematic reviews and meta-analyses have been conducted so far including an analysis to investigate the difference between ethnic groups in glucose lowering efficacy of DPP-4 inhibitors. Kim et al. [15] reported improved glucose lowering efficacy in Asians compared with other ethnic groups, and Park et al. [16]

reported the difference between Japanese and non-Japanese patients in efficacy and safety. However there were several limitations in these analyses. First, the number of articles which deal with clinical studies in the Asian or Japanese population is few compared to that in the Western population. Second, seven Japanese studies referred in the above mentioned articles showed a greater reduction in HbA1c such as -1.67% , which is far greater than the generally recognized value such as -0.6% to -0.9% . Third, HbA1c as it had been expressed in Japan used the Japan Diabetes Society (JDS) system values and not the National Glycohemoglobin Standardization Program (NGSP) values. From April 1, 2013, the HbA1c results are reported using only the NGSP values and there is a possibility that meta-analysis in the above mentioned articles were conducted in different HbA1c values.

As for the SGLT-2 inhibitor, which appeared on the market most recently, a pooled analysis was conducted to assess the efficacy and safety in patients of different ethnicities, Hispanic/Latino and non-Hispanic/Latino [17]. But systematic reviews and meta-analysis to investigate the difference in efficacy and safety of SGLT-2 inhibitors between different ethnic groups considering the pathophysiology such as Asian and non-Asian or Japanese and non-Japanese like DPP-4 inhibitor has not been conducted so far. The reason might be its insulin independent mechanism of action.

The placebo phenomenon in chronic pain or neuropsychiatric disorder has been widely recognized in clinical research because high placebo responses were reported to be the cause of failure of clinical studies. In type 2 diabetes area, however, placebo effect has rarely been discussed.

Against this background, we performed a systematic review and meta-analysis of randomized controlled studies for type 2 diabetes and assessed the correlation

between the glucose-lowering efficacy of oral hypoglycemic agents and potential factors related to clinical study design and subject demographic, and conduct including the ethnicity of study subjects and study regions. As oral hypoglycemic agents, we selected two classes of mechanism of actions, a DPP-4 inhibitor, which has an insulin secretion mechanism (Research 1), and a SGLT-2 inhibitor, which has an insulin independent mechanism (Research 2). Additionally, we conducted an analysis to identify factors contributing to placebo effect in clinical studies for type 2 diabetes. Based on these analyses, we discuss the points to be considered in future clinical development of oral hypoglycemic drugs.

2. Research 1 (DPP-4 inhibitor)

2.1. Objectives

DPP-4 inhibitors are a class of oral hypoglycemic agents that inhibit the enzyme DPP-4. The enzyme breaks down the incretin gastrointestinal hormones glucagon-like peptide-1 and glucose-dependent insulintropic polypeptide, which are released in response to a meal and stimulate glucose-dependent insulin secretion [12, 13]. DPP-4 inhibitors are now in widespread use in clinical practice including East Asian countries and are rapidly becoming first-line therapy [14].

The pathophysiological differences may have an impact on the therapeutic approach and several systematic reviews and meta-analyses have been conducted including an analysis to investigate the difference between ethnic groups in glucose lowering efficacy of DPP-4 inhibitors. Kim et al. [15] reported improved glucose lowering efficacy in Asians compared with other ethnic groups, and Park et al. [16] reported the difference between Japanese and non-Japanese patients in efficacy and safety. Others [18-20] have analyzed the correlation between glucose-lowering efficacy and baseline characteristics without including ethnic information. However, no consistent result has been found to date.

Thus, we performed a systematic review and meta-analysis of randomized controlled studies and conducted meta-regression analyses to assess the correlation between the glucose-lowering efficacy of DPP-4 inhibitors including alogliptin, linagliptin, sitagliptin, saxagliptin and vildagliptin in patients with type 2 diabetes and potential factors including ethnicity of study subjects and study regions; in particular, Japanese were dealt separately from other Asians.

2.2. Method

The study conduct and results were reported in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) statement [21].

2.2.1. Data sources and searches

A systematic search was performed to identify potentially relevant studies in Medline, Embase, Cochrane Central Register, Japan Medical Abstracts Society and ClinicalTrials.gov databases up to October 4, 2016. In addition, the search included the reviews of approved drugs from the US Food and Drug Administration (www.fda.gov), European Medicines Agency (www.ema.europa.eu) and Japanese Pharmaceuticals and Medical Devices Agency (www.pmda.go.jp). An extensive search for alogliptin, linagliptin, sitagliptin, saxagliptin and vildagliptin was performed.

Eligibility criteria for study selection included the following: (1) randomized controlled studies comparing a DPP-4 inhibitor with a placebo as either monotherapy or combination therapy with other oral glucose-lowering drugs in patients with type 2 diabetes, (2) studies with treatment duration of at least 12 weeks, (3) studies with information on HbA1c values of placebo-adjusted change from baseline, and (4) studies published or described in English or Japanese. Eligibility was assessed independently by two authors (K.F. and M.K.), and disagreements were resolved by consensus with another author (M.N.).

Exclusion criteria included the following: (1) studies with duration of less than 12 weeks because of inadequate time to assess HbA1c change, (2) duplicate and extended studies from original studies, and (3) studies lacking relevant information.

2.2.2. Data extraction

Two authors (K.F. and M.K.) independently conducted data extraction. Any disagreements were resolved by consensus with another author (M.N.). The following data were extracted from the eligible studies: study design (monotherapy or oral combination therapy), name of the first author, year of publication, country where the study was conducted, treatment regimen, study duration, number of participants, percentage of male, means of age, duration of diabetes, baseline HbA1c values, baseline fasting plasma glucose, baseline body mass index (BMI), baseline body weight, and percentages of Asians and Japanese subjects. The primary outcome was the effect of DPP-4 inhibitors on HbA1c change from baseline compared with placebo at each study's primary endpoint.

For dose-ranging studies, only data from currently approved medication doses were extracted. In the absence of such data, equivalent amounts of daily doses were used. If a study had two or three comparisons (one monotherapy arm and one or two combination therapy arms), each comparison was treated separately. When there was a comparison for different administration times in a day (morning vs. evening), both data were used. Some data not available in the original paper were subsequently extracted using information from a full report available in the study registry, ClinicalTrials.gov, or from reviews of approved drugs by regulatory agencies.

For ethnicity information, we followed the classification of 'Asian' used by the authors of each study. If a study did not disclose ethnic information but was conducted in Asian countries with a relatively homogeneous population, e.g., Korea, China, Japan or Taiwan, we assumed the study participants were Asian. In addition, if a study did not disclose ethnic information but was conducted only in Japan, we assumed the study

participants were Japanese.

In order to further explore potential differences in HbA1c-lowering efficacy among the ethnicity and study regions, we created three classifications, Type A, B and C. Type A and Type B focused on ethnicity of study subjects, and Type C focused on study regions. In Type A classification, Japanese was included in Asian, and we categorized the studies either as a “study in Asian subjects” or a “study in non-Asian subjects” based on whether the percentage of Asians was larger than or equal to 50% or not. In Type B, Japanese was dealt separately from Asian, and we divided the studies into three groups: a “study in Japanese subjects” that enrolled only Japanese, a “study in Asian subjects” that excluded the study in Japanese subjects from study in Asian subjects of Type A, and a “study in non-Asian subjects.” In Type C, we divided the studies into three groups: a “study conducted in Japan” that conducted only in Japan, an “study conducted in Asia” conducted only in Asia and excluding those conducted only in Japan, and a “study conducted multinationally.”

For HbA1c value, when the Japan Diabetes Society (JDS) values, HbA1c (JDS), were used in Japanese studies [22], then they were calculated into National Glycohemoglobin Standardization Program (NGSP) equivalent value (%) using the formula $HbA1c (\%) = HbA1c (JDS) (\%) + 0.4\%$ and used in the present study.

2.2.3. Study quality and risk of bias assessment

The Jadad Scale was adopted to assess the quality of all included studies. We evaluated the included studies regarding randomization, concealment of allocation, double-blinding, withdrawals and dropouts. To assess publication bias, we used the funnel plot and Egger’s test.

2.2.4. Statistical analysis

Weighted mean difference (WMD) and 95% confidence interval for HbA1c change from baseline in DPP-4 inhibitors compared with placebo was calculated. Heterogeneity between studies was assessed by using Q statistic and I^2 statistic. A P -value of Q statistic <0.10 and a value of 50% or greater of I^2 statistic was considered significant. A random-effect model was used because significant heterogeneity was reported in several meta-analyses [15,16]. Meta-regression was applied to investigate potential factors affecting the placebo adjusted HbA1c lowering efficacy. The pre-specified factors such as study duration, percentage of male, age, duration of diabetes, baseline HbA1c values, fasting plasma glucose, BMI, body weight, percentage of Asians and additional categorizations regarding ethnicity of study subjects and study regions were used in the meta-regression.

Univariate meta-regression was performed first to identify potential factors affecting the HbA1c-lowering efficacy of the DPP-4 inhibitor. Negative coefficient indicates efficacy is in favor of the factor. A significant association was defined as $P <0.2$ in univariate meta-regression. Factors were further analyzed by multivariate meta-regression. When correlation between explanatory variables was identified, one variable was selected to be included in the multivariate meta-regression. A statistically significant difference was defined as $P <0.05$ in the multivariate meta-regression. Analyses were performed using StatDirect version 2.7.9 (StatDirect Ltd., Altrincham, Cheshire, UK) and R software, version 3.2.0 [23].

2.3. Result

2.3.1. Literature search

We identified 2,721 potentially relevant articles from a literature search. Based

on the review of the abstracts, 72 studies met all of the inclusion criteria. Further 270 studies were identified in the ClinicalTrials.gov registry and among them 7 unpublished clinical studies were included in the analysis (Fig. 1). A total of 79 studies (Appendix: reference 1-79) with 91 arms and 25,095 patients were used for the analysis (Table 1). Regarding the factors related to ethnicity and regions, in Type B, 17 were studies in Asian subjects, 18 were studies in Japanese subjects and 41 were studies in non-Asian subjects. As for Type C, 6 were studies conducted in Asia, 18 were studies conducted in Japan and 55 were studies conducted multinationally (Table 2).

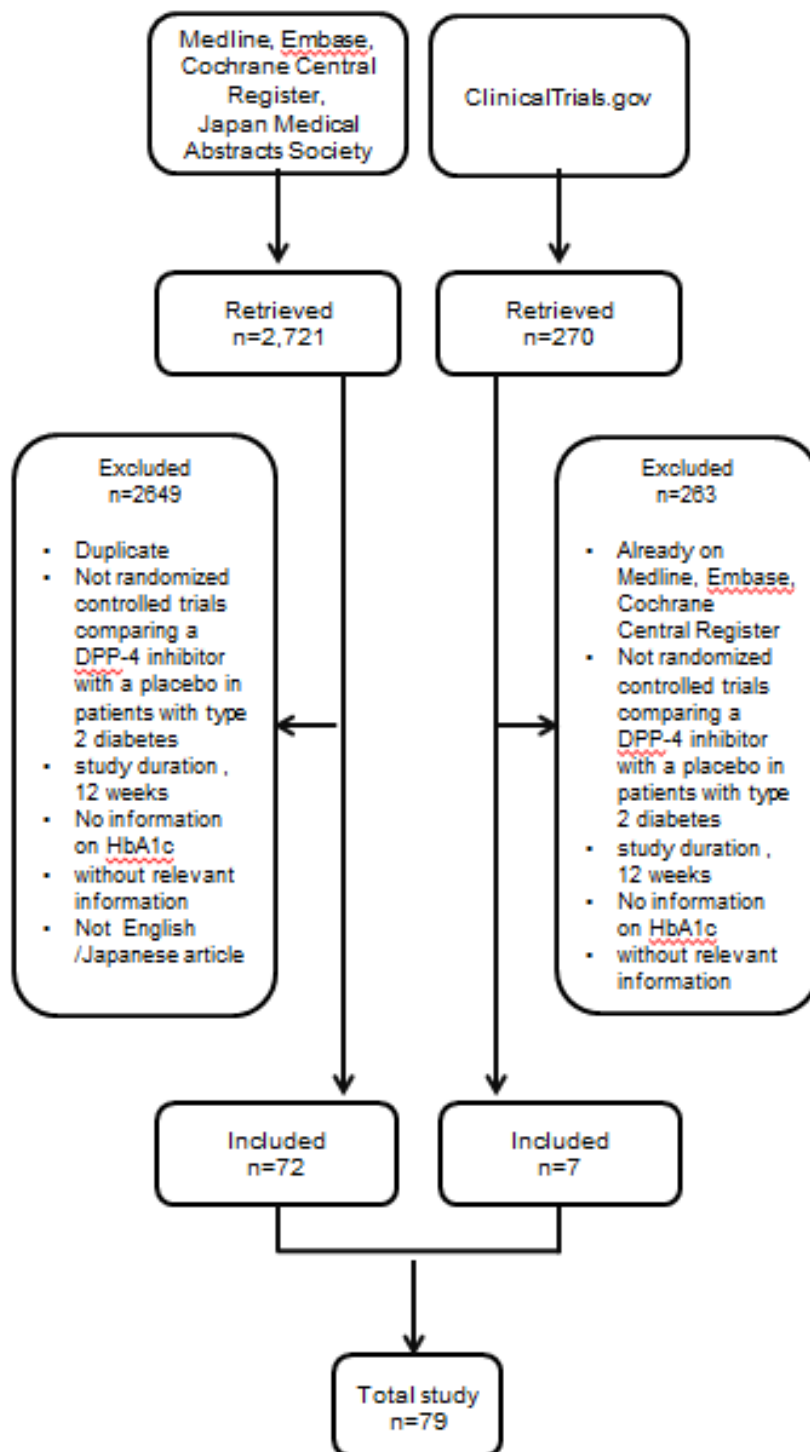


Figure 1 Flow diagram of study selection

Table 1 Studies included in the analysis

Source	Number of participants	Treatment duration (weeks)	DPP-4 inhibitors	background antidiabetic therapy	Age (years)	Baseline HbA1c (%)
DeFronzo, 2008 ¹	196	26	Alogliptin 25 mg QD	none	53.4	8.0
Nauck, 2009 ²	314	26	Alogliptin 25 mg QD	metformin	54.7	7.9
Pratley, 2009 ³	296	26	Alogliptin 25 mg QD	pioglitazone	55.3	8.0
Pratley, 2009 ⁴	297	26	Alogliptin 25 mg QD	glyburide	56.7	8.1
Seino, 2011 ⁵	155	12	Alogliptin 25 mg QD	none	59.3	7.9
Kaku, 2011 ⁶	228	12	Alogliptin 25 mg QD	pioglitazone	59.7	7.9
Seino, 2011 ⁷	154	12	Alogliptin 25 mg QD	voglibose	62.6	8.0
Seino, 2012 ⁸	207	12	Alogliptin 25 mg QD	glimepiride	60.0	8.6
Seino, 2012 ⁹	196	12	Alogliptin 25 mg QD	metformin	52.2	8.0
Pan, 2016 ¹⁰ _1 ^a	185	16	Alogliptin 25 mg QD	none	52.4	8.0
Pan, 2016 ¹⁰ _2 ^b	197	16	Alogliptin 25 mg QD	metformin	53.1	8.0
Pan, 2016 ¹⁰ _3 ^c	124	16	Alogliptin 25 mg QD	pioglitazone (with or without metformin)	52.2	8.0
Forst, 2010 ¹¹	137	12	Linagliptin 5 mg QD	metformin	59.9	8.4
Del Prato, 2011 ¹²	503	24	Linagliptin 5 mg QD	none	55.7	8.0
Taskinen, 2011 ¹³	700	24	Linagliptin 5 mg QD	metformin	56.5	8.1
Owens, 2011 ¹⁴	1055	24	Linagliptin 5 mg QD	metformin and sulfonylurea	58.1	8.1
Lewin, 2012 ¹⁵	245	18	Linagliptin 5 mg QD	sulfonylurea	56.9	8.6
Kawamori, 2012 ¹⁶	239	12	Linagliptin 5 mg QD	none	60.1	8.0

Source	Number of participants	Treatment duration (weeks)	DPP-4 inhibitors	background antidiabetic therapy	Age (years)	Baseline HbA1c (%)
Ross, 2012 ¹⁷ _1 ^a	268	12	Linagliptin 5 mg QD	metformin	58.6	8.0
Ross, 2012 ¹⁷ _2 ^b	267	12	Linagliptin 2.5 mg BID	metformin	58.9	8.0
Haak, 2012 ¹⁸	214	24	Linagliptin 5 mg QD	none	56.0	8.7
Barnett, 2012 ¹⁹	227	18	Linagliptin 5 mg QD	none	56.5	8.1
Bajaj, 2014 ²⁰	278	24	Linagliptin 5 mg QD	metformin and pioglitazone	53.8	8.4
Chen, 2015 ²¹	301	24	Linagliptin 5 mg QD	none	54.4	8.0
Wang, 2016 ²²	305	24	Linagliptin 5 mg QD	metformin	55.6	8.0
Rosenstock, 2008 ²³	114	12	Saxagliptin 5 mg QD	none	54.6	8.0
Rosenstock, 2009 ²⁴	201	24	Saxagliptin 5 mg QD	none	53.9	8.0
DeFronzo, 2009 ²⁵	370	24	Saxagliptin 5 mg QD	metformin	54.7	8.1
Hollander, 2009 ²⁶	370	24	Saxagliptin 5 mg QD	thiazolidinedione	53.6	8.3
Chacra, 2009 ²⁷	520	24	Saxagliptin 5 mg QD	glyburide	55.0	8.4
Yang, 2011 ²⁸	570	24	Saxagliptin 5 mg QD	metformin	54.1	7.9
Pan, 2012 ²⁹	568	24	Saxagliptin 5 mg QD	none	51.4	8.2
Frederich, 2012 ³⁰ _1 ^a	148	24	Saxagliptin 5 mg QD AM	none	55.1	7.9
Frederich, 2012 ³⁰ _2 ^b	146	24	Saxagliptin 5 mg QD PM	none	55.3	7.9
White, 2014 ³¹	160	12	Saxagliptin 2.5 mg BID	metformin	55.4	7.9
Moses, 2014 ³²	257	24	Saxagliptin 5 mg QD	metformin and sulfonylurea	57.0	8.3
Seino, 2014 ³³ _1 ^a	168	12	Saxagliptin 5 mg QD	none	58.6	8.5
Seino, 2014 ³³ _2 ^b	187	24	Saxagliptin 5 mg QD	none	57.9	8.0

Source	Number of participants	Treatment duration (weeks)	DPP-4 inhibitors	background antidiabetic therapy	Age (years)	Baseline HbA1c (%)
Matthaei, 2015 ³⁴	315	24	Saxagliptin 5 mg QD	dapagliflozin and metformin	54.6	7.9
NCT00918879 ³⁵	213	24	Saxagliptin 5 mg QD	none	48.7	8.3
Ristic, 2005 ³⁶	121	12	Vildagliptin 100 mg QD	none	55.4	7.7
Dejager, 2007 ³⁷ _1 ^a	317	24	Vildagliptin 100 mg QD	none	52.9	8.4
Dejager, 2007 ³⁷ _2 ^b	312	24	Vildagliptin 50 mg BID	none	52.5	8.5
Pi-Sunyer, 2007 ³⁸ _1 ^a	183	24	Vildagliptin 100 mg QD	none	52.0	8.4
Pi-Sunyer, 2007 ³⁸ _2 ^b	175	24	Vildagliptin 50 mg BID	none	51.1	8.5
Bosi, 2007 ³⁹	367	24	Vildagliptin 100 mg QD	metformin	54.2	8.4
Garber, 2007 ⁴⁰	316	24	Vildagliptin 100 mg QD	pioglitazone	54.4	8.7
Garber, 2008 ⁴¹	345	24	Vildagliptin 100 mg QD	glimepiride	58.0	8.5
Kikuchi, 2009 ⁴²	148	12	Vildagliptin 50 mg BID	none	59.6	7.8
Kikuchi, 2010 ⁴³	202	12	Vildagliptin 50 mg BID	glimepiride	59.7	8.3
Kikuchi, 2010 ⁴⁴ _1 ^a	124	12	Vildagliptin 50 mg BID	none	59.5	7.7
Kikuchi, 2010 ⁴⁴ _2 ^b	118	12	Vildagliptin 100 mg QD	none	61.1	7.7
Pan, 2012 ⁴⁵	290	24	Vildagliptin 50 mg BID	metformin	54.3	8.0
Odawara, 2014 ⁴⁶	139	12	Vildagliptin 50 mg BID	metformin	58.1	8.0
Raz, 2006 ⁴⁷	315	18	Sitagliptin 100 mg QD	none	54.8	8.0
Aschner, 2006 ⁴⁸	491	24	Sitagliptin 100 mg QD	none	53.9	8.0
Charbonnel, 2006 ⁴⁹	701	24	Sitagliptin 100 mg QD	metformin	54.5	8.0
Rosenstock, 2006 ⁵⁰	353	24	Sitagliptin 100 mg QD	pioglitazone	56.3	8.0

Source	Number of participants	Treatment duration (weeks)	DPP-4 inhibitors	background antidiabetic therapy	Age (years)	Baseline HbA1c (%)
Hanefeld, 2007 ⁵¹ _1 ^a	221	12	Sitagliptin 100 mg QD	none	55.9	7.7
Hanefeld, 2007 ⁵¹ _2 ^b	222	12	Sitagliptin 50 mg BID	none	55.6	7.7
Hanefeld, 2007 ⁵¹ _3 ^c	223	12	Sitagliptin 50 mg QD	none	55.9	7.6
Goldstein, 2007 ⁵²	355	24	Sitagliptin 100 mg QD	none	53.4	8.8
Hermansen, 2007 ⁵³ _1 ^a	212	24	Sitagliptin 100 mg QD	glimepiride	54.8	8.4
Hermansen, 2007 ⁵³ _2 ^b	229	24	Sitagliptin 100 mg QD	glimepiride and metformin	57.1	8.3
Scott, 2007 ⁵⁴	249	12	Sitagliptin 50 mg BID	none	55.2	7.9
Raz, 2008 ⁵⁵	190	30	Sitagliptin 100 mg QD	metformin	54.8	9.2
Scott, 2008 ⁵⁶	186	18	Sitagliptin 100 mg QD	metformin	55.2	7.8
Nonaka, 2008 ⁵⁷	152	12	Sitagliptin 100 mg QD	none	55.3	8.0
Mohan, 2009 ⁵⁸	530	18	Sitagliptin 100 mg QD	none	50.9	8.7
Iwamoto, 2010 ⁵⁹ _1 ^a	143	12	Sitagliptin 100 mg QD	none	59.3	8.1
Iwamoto, 2010 ⁵⁹ _2 ^b	145	12	Sitagliptin 50 mg QD	none	60.2	8.1
Kashiwagi, 2011 ⁶⁰	134	12	Sitagliptin 50 mg QD	pioglitazone	58.4	8.1
Barzilai, 2011 ⁶¹	206	24	Sitagliptin 50 or 100 mg QD	none	71.9	7.8
Tajima, 2011 ⁶²	138	12	Sitagliptin 50 mg QD	glimepiride	60.7	9.1
Yang, 2012 ⁶³	395	24	Sitagliptin 100 mg QD	metformin	54.6	8.5
Nicolle, 2012 ⁶⁴	130	12	Sitagliptin 100 mg QD	metformin	52.5	7.7
Fonseca, 2013 ⁶⁵	313	26	Sitagliptin 100 mg QD	metformin and pioglitazone	56.0	8.8

Source	Number of participants	Treatment duration (weeks)	DPP-4 inhibitors	background antidiabetic therapy	Age (years)	Baseline HbA1c (%)
Dobs, 2013 ⁶⁶	262	18	Sitagliptin 100 mg QD	metformin and rosiglitazone	54.5	8.8
Roden, 2013 ⁶⁷	451	24	Sitagliptin 100 mg QD	none	55.0	7.9
Lavalle-González, 2013 ⁶⁸	549	26	Sitagliptin 100 mg QD	metformin	55.4	7.9
Tajima, 2013 ⁶⁹	133	12	Sitagliptin 50 mg QD	voglibose	60.5	7.9
Kadowaki, 2013 ⁷⁰	149	12	Sitagliptin 50 mg QD	metformin	58.4	7.9
Amin, 2015 ⁷¹	112	12	Sitagliptin 100 mg QD	metformin	47.5	8.0
Ji, 2016 ⁷²	247	24	Sitagliptin 100 mg QD	none	52.7	8.9
Moses, 2016 ⁷³	422	24	Sitagliptin 100 mg QD	sulfonylurea and metformin	54.9	8.4
NCT01177384 ⁷⁴	381	24	Sitagliptin 100 mg QD	acarbose	57.1	8.1
NCT01289990 ⁷⁵	451	52	Sitagliptin 100 mg QD	none	55.0	7.9
NCT01338870 ⁷⁶	100	12	Sitagliptin 100 mg QD	metformin	55.7	8.1
NCT01336738 ⁷⁷	107	12	Sitagliptin 100 mg QD	metformin	56.7	8.3
NCT01590771 ⁷⁸	498	24	Sitagliptin 100 mg QD	sulfonylurea with or without metformin	57.0	8.6
NCT01703221 ⁷⁹	248	24	Sitagliptin 50 mg QD	none	60.3	NR

QD once-daily dosing, *BID* twice-daily dosing, *NR* not reported

^a 1 represents the first pair of the article

^b 2 represents the second pair of the article

^c 3 represents the third pair of the article

Table 2 Three classifications for studies

Type A	Study in non-Asian subjects 41 studies	Study in Asian subjects 35 studies	
Type B	Study in non-Asian subjects 41 studies	Study in Asian subjects (excluding the Japanese) 17 studies	Study in Japanese subjects ^a 18 studies
Type C	Study conducted multinationally 55 studies	Study conducted in Asia (excluding Japan) 6 studies	Study conducted in Japan ^a 18 studies

^a Studies classified as "Study in Japanese subjects" (Type B) and "Study conducted in Japan" (Type C) were the same.

2.3.2. Quality of included studies and publication bias

The Jaded score was 3.74, and we determined the quality of included studies was high. The funnel plot of WMD for HbA1c and standard error in 79 studies is shown in Fig. 2. Funnel plot asymmetry was assessed using Egger's test and it indicated that there was no publication bias ($p = 0.067$). The I^2 statistic test for heterogeneity was 69.7%, and the value of Q statistic was significant.

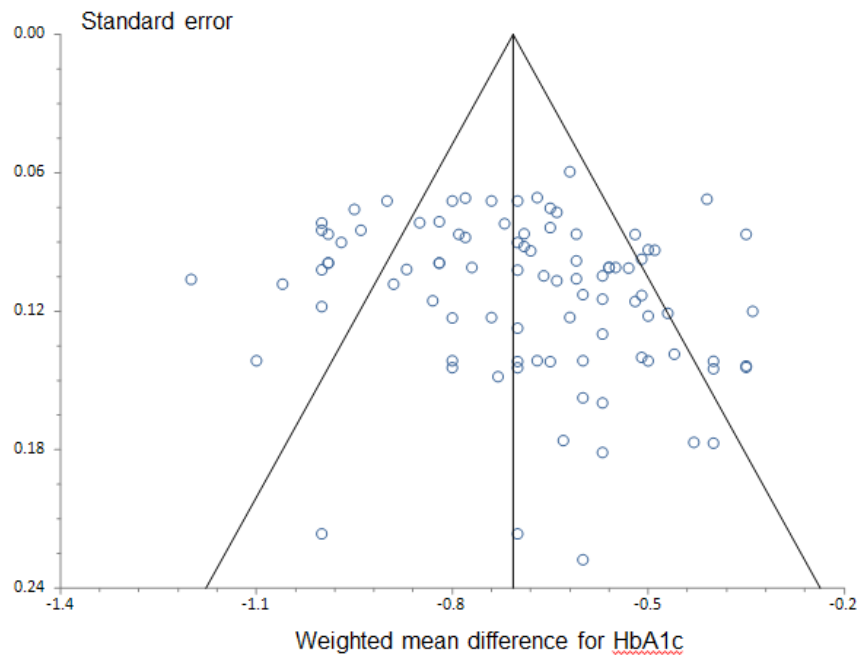


Figure 2 Funnel plot of weighted mean difference for HbA1c and standard error

2.3.3. Primary outcome

HbA1c data for 91 arms from 79 studies were pooled. The difference of HbA1c between the DPP-4 inhibitor group and placebo group was -0.695% (95% CI -0.734 to -0.656).

2.3.4. Univariate meta-regression

Univariate meta-regression showed relations between the WMD in the change of HbA1c from baseline value and several factors such as study duration (coefficient=0.009, $p = 0.001$), percentage of male (-0.009 , $p < 0.001$), age (-0.024 , $p < 0.001$), duration of diabetes (-0.034 , $p = 0.001$), baseline HbA1c values (-0.087 , $p = 0.162$), BMI (0.029 , $p < 0.001$), body weight (0.009 , $p < 0.001$) and percentage of Asians (-0.002 , $p < 0.001$). Additionally, studies in Asian subjects in Type A (-0.129 , $p = 0.001$), studies in Japanese subjects in Type B (-0.254 , $p < 0.001$), and studies conducted in Japan in Type C (-0.263 , $p < 0.001$) showed relations in each of the analyses. Because BMI and body weight correlated with the percentage of Asians, we excluded BMI and body weight from the multivariate meta-regression, and used the three classifications (Type A, B and C).

2.3.5. Multivariate meta-regression

The factors significantly associated with the WMD in the change of HbA1c from baseline using univariate meta-regression were further analyzed by multivariate meta-regression (Table 3). Baseline HbA1c was identified as a common influencing factor in all the analyses using the three types of classifications. Also, study in Japanese subjects in Type B and study conducted in Japan in Type C were identified as influencing factors.

Table 3 Results of multivariate meta-regression analysis

Factors	Estimate	95% CI	P-value
a. Multivariate meta-regression with Type A			
Study duration (week)	0.007	(-0.003, 0.016)	0.166
Male (%)	-0.003	(-0.010, 0.004)	0.392
Age (years)	-0.011	(-0.027, 0.004)	0.138
Duration of diabetes (years)	-0.012	(-0.034, 0.011)	0.302
Baseline HbA1c (%)	-0.169	(-0.300, -0.037)	0.012
Study in Asian subjects	-0.078	(-0.176, 0.019)	0.116
b. Multivariate meta-regression with Type B			
Study duration (week)	0.003	(-0.006, 0.013)	0.502
Male (%)	0.001	(-0.006, 0.008)	0.684
Age (years)	-0.002	(-0.018, 0.013)	0.777
Duration of diabetes (years)	-0.007	(-0.029, 0.014)	0.499
Baseline HbA1c (%)	-0.157	(-0.281, -0.032)	0.014
Study in Asian subjects	0.018	(-0.096, 0.132)	0.755
Study in Japanese subjects	- 0.242	(-0.389, -0.095)	0.001
c. Multivariate meta-regression with Type C			
Study duration (week)	0.003	(-0.007, 0.012)	0.588
Male (%)	0.001	(-0.006, 0.008)	0.710
Age (years)	-0.004	(-0.019, 0.011)	0.605
Duration of diabetes (years)	-0.008	(-0.029, 0.014)	0.484
Baseline HbA1c (%)	-0.157	(-0.281, -0.032)	0.013
Study conducted in Asia	0.011	(-0.129, 0.152)	0.873
Study conducted in Japan	- 0.243	(-0.391, -0.096)	0.001

Type A, the studies were divided into 2 groups; a “study in Asian subjects” or a “study in non-Asian subjects”

Type B, the studies were divided into 3 groups; a “study in Asian subjects”, a “study in Japanese subjects” or a “study in non-Asian subjects.”

Type C, the studies were divided into three groups: a “study conducted in Asia”, a “study conducted in Japan” or a “study conducted multinationally.”

HbA1c hemoglobin A1c

2.4. Discussion

We performed meta-regression analyses to identify the factors influencing the glucose-lowering efficacy of DPP-4 inhibitors by analyzing 79 studies with 91 arms and 25,095 patients. We found WMD in the placebo-adjusted change of HbA1c from baseline was -0.695% , which is consistent with the previously reported meta-analyses ranging between -0.6% and -0.72% .

Using univariate meta-regression, factors such as study duration, percentage of male, age, duration of diabetes, baseline HbA1c values, BMI, body weight and percentage of Asians were identified to be correlated with WMD in the change of HbA1c from baseline. The results of study duration, percentage of male, age, duration of diabetes, BMI, and percentage of Asians were in agreement with the analysis conducted by Kim et al [15]. Additionally, in the present study, relations were shown in studies in Asian subjects in the classification Type A, studies in Japanese subjects in Type B and studies conducted in Japan in Type C. Based on these results, DPP-4 inhibitors exhibit a better glucose-lowering efficacy with shorter study duration, in older patients, with longer duration of diabetes, with higher baseline HbA1c, with lower BMI and body weight, and with a higher percentage of Asians. A better glucose lowering efficacy was also shown in studies in Asian subjects, studies in Japanese subjects and studies conducted in Japan.

Multivariate meta-regression showed a better glucose lowering efficacy with high baseline HbA1c, suggesting baseline HbA1c is a robust factor. Moreover, studies in Japanese subjects in Type B and studies conducted in Japan in Type C were factors influencing the glucose-lowering efficacy of DPP-4 inhibitors.

Bloomgarden et al. [24] reported that the baseline glycemic status of patients

recruited into clinical studies strongly influences the reduction of HbA1c following pharmacologic intervention with the five major anti-hyperglycemic oral agent classes. The authors concluded that clinical investigators and study sponsors should consider these findings when testing the effectiveness of new anti-hyperglycemic agents. A higher baseline HbA1c would be a predictor of a greater HbA1c reduction with the use of DPP-4 inhibitors [20, 25]. However, the HbA1c value used by Esposito et al. [20] and Deacon [25] in their analysis was different from our study; they used the absolute HbA1c reduction without taking into account the placebo effect. In the present study, we found that a higher baseline HbA1c is a factor of greater efficacy of DPP-4 inhibitors on WMD in the change of HbA1c from baseline. For clinical studies, the comparison with placebo is generally required. Therefore, in evaluating the efficacy of DPP-4 inhibitors in clinical development, it is important to provide the information on factors that influence the efficacy of DPP-4 inhibitors. Our finding suggests that paying attention to baseline HbA1c in planning and conducting of clinical studies, and comparing the data with other studies is important.

In the present study, we used three classifications regarding ethnicity of study subjects and study regions to explore potential differences among them. In the classification Type A, Japan was included in Asia. In Type B and C, we distinguished between Asia and Japan focusing on ethnicity in Type B and on regions in Type C. Studies in Japanese subjects and studies conducted in Japan are factors affecting the efficacy of DPP-4 inhibitors. However, studies in Asian subjects in Type B and studies conducted in Asia in Type C are not associated. This finding differs from those by Kim et al. [15], who reported that the DPP-4 inhibitors exhibit a better glucose-lowering efficacy in Asians than in other ethnic groups. Our finding indicates that differences in

the contribution of the insulin secretory defect and insulin resistance in the pathophysiology of type 2 diabetes between Asians and non-Asians does not significantly affect the response to DPP-4 inhibitors and does not explain the mechanism underlying the difference in the response.

The differences in the efficacy of DPP-4 inhibitors between Japanese and non-Japanese patients have been reported. Interestingly, Park et al. [16] suggested that the difference in characteristics such as age, duration of diabetes, baseline HbA1c and BMI between Japanese and non-Japanese patients could provide some explanation for the discrepancy in clinical studies. In their analysis, however, they did not clearly mention whether they took into consideration the difference of HbA1c values between JDS and NGSP. In our study, if an investigation in Japan did not mention the value of HbA1c (%), we checked the documents, review reports or other sources in order to use HbA1c (NGSP) equivalent values (%) for the analysis. The mean baseline HbA1c values were not significantly different among studies in non-Asian subjects, studies in Asian subjects and studies in Japanese subjects (data not shown), and the baseline HbA1c was an independent factor associated with the efficacy of DPP-4 inhibitors in our study.

When we looked at the placebo response in the classification Type B and Type C, the trend of the response was different among the three groups. Placebo response was increased in studies in Japanese subjects and studies conducted in Japan, whereas it decreased in studies in Asian subjects and studies conducted in Asia; in studies in non-Asian subjects and studies conducted multinationally it showed minimal change from baseline (data not shown). Further investigation is needed for better understanding the reason for the difference. In the present study, the factors analyzed in the univariate

meta-regression do not clearly explain the difference in efficacy of the DPP4-inhibitor in studies in Japanese subjects/studies conducted in Japan.

Our analysis indicates that studies in Japanese subjects /studies conducted in Japan as well as the baseline HbA1c values are associated with the efficacy of DPP-4 inhibitors. This result should be taken into account when planning and conducting clinical studies. Further investigation is needed to understand the association of the differences in the contribution of insulin secretory defect and insulin resistance in the pathophysiology of type 2 diabetes among ethnicity with the glucose-lowering efficacy of DPP-4 inhibitors.

3. Research 2 (SGLT-2 inhibitor)

3.1. Objectives

Most recently approved oral hypoglycemic drugs with a new mechanism is SGLT2 inhibitors. Inhibition of SGLT-2 decreases reabsorption of glucose in proximal renal tubular cells, leading to an increase in urinary glucose excretion and a reduction in plasma glucose levels. SGLT-2 inhibitors entered the Japanese market in 2014. Canagliflozin, dapagliflozin, and empagliflozin have been approved for use in the United States, the European Union and Japan, while ipragliflozin, luseogliflozin, and tofogliflozin have been approved for use only in Japan.

A pooled analysis was conducted to assess the efficacy and safety in patients of different ethnicities by Davidson et al. [17]. They reported that the SGLT2 inhibitor canagliflozin is equally effective and generally well-tolerated in both Hispanic/Latino and non-Hispanic/Latino. They focused on this ethnicity because Hispanic/Latino patients are one of the large populations with type 2 diabetes in the United States. But systematic reviews and meta-analysis to investigate the difference in efficacy and safety of SGLT-2 inhibitors between different ethnic groups considering the pathophysiology such as Asian and non-Asian or Japanese and non-Japanese like DPP-4 inhibitor has not been conducted so far. The reason might be its insulin independent mechanism of action, neither insulin secretion nor sensitization.

Thus, we performed a systematic review and meta-analysis of randomized controlled studies and conducted meta-regression analyses to assess the correlation between the glucose-lowering efficacy of SGLT-2 inhibitors including dapagliflozin, canagliflozin and empagliflozin in patients with type 2 diabetes and potential factors including ethnicity of study subjects and study regions; in particular, Japanese were

dealt separately from other Asians.

3.2. Method

The study conduct and results were reported in accordance with the PRISMA statement [21].

3.2.1. Data sources and searches

A systematic search was performed to identify potentially relevant studies in Medline, Embase, Cochrane Central Register, Japan Medical Abstracts Society and ClinicalTrials.gov databases up to December 31, 2015. In addition, the search included the reviews of approved drugs from the US Food and Drug Administration (www.fda.gov), European Medicines Agency (www.ema.europa.eu) and Japanese Pharmaceuticals and Medical Devices Agency (www.pmda.go.jp). An extensive search for dapagliflozin, canagliflozin and empagliflozin was performed.

Eligibility criteria for study selection included the following: (1) randomized controlled studies comparing aSGLT-2 inhibitor with a placebo as either monotherapy or combination therapy with other oral glucose-lowering drugs in patients with type 2 diabetes, (2) studies with treatment duration of at least 12 weeks, (3) studies with information on HbA1c values of placebo-adjusted change from baseline, and (4) studies published or described in English or Japanese. Eligibility was assessed independently by two authors (K.F. and M.K.), and disagreements were resolved by consensus with another author (M.N.).

Exclusion criteria included the following: (1) studies with duration of less than 12 weeks because of inadequate time to assess HbA1c change, (2) duplicate and

extended studies from original studies, and (3) studies lacking relevant information.

3.2.2. Data extraction

Two authors (K.F. and M.K.) independently conducted data extraction. Any disagreements were resolved by consensus with another author (M.N.). The following data were extracted from the eligible studies: study design (monotherapy or oral combination therapy), name of the first author, year of publication, country where the study was conducted, treatment regimen, study duration, number of participants, percentage of male, means of age, duration of diabetes, baseline HbA1c values, baseline fasting plasma glucose, BMI, baseline body weight, and percentages of Asians and Japanese subjects. The primary outcome was the effect of SGLT-2 inhibitors on HbA1c change from baseline compared with placebo at each study's primary endpoint.

For dose-ranging studies, only data from currently approved medication doses were extracted. In the absence of such data, equivalent amounts of daily doses were used. If a study had two or three comparisons (one monotherapy arm and one or two combination therapy arms), each comparison was treated separately. When there was a comparison for different administration times in a day (morning vs. evening), both data were used. Some data not available in the original paper were subsequently extracted using information from a full report available in the study registry, ClinicalTrials.gov, or from reviews of approved drugs by regulatory agencies.

For ethnicity information, we followed the classification of 'Asian' used by the authors of each study. If a study did not disclose ethnic information but was conducted in Asian countries with a relatively homogeneous population, e.g., Korea, China, Japan or Taiwan, we assumed the study participants were Asian. In addition, if a study did not

disclose ethnic information but was conducted only in Japan, we assumed the study participants were Japanese.

In order to further explore potential differences in HbA1c-lowering efficacy among the ethnicity and study regions, we created three classifications, Type A, B and C. Type A and Type B focused on ethnicity of study subjects, and Type C focused on study regions. In Type A classification, Japanese was included in Asian, and we categorized the studies either as a “study in Asian subjects” or a “study in non-Asian subjects” based on whether the percentage of Asians was larger than or equal to 50% or not. In Type B, Japanese was dealt separately from Asian, and we divided the studies into three groups: a “study in Japanese subjects” that enrolled only Japanese, a “study in Asian subjects” that excluded the study in Japanese subjects from study in Asian subjects of Type A, and a “study in non-Asian subjects.” In Type C, we divided the studies into three groups: a “study conducted in Japan” that conducted only in Japan, an “study conducted in Asia” conducted only in Asia and excluding those conducted only in Japan, and a “study conducted multinationally.”

For HbA1c value, when the JDS values, HbA1c (JDS), were used in Japanese studies [22], then they were calculated into NGSP equivalent value (%) using the formula $\text{HbA1c (\%)} = \text{HbA1c (JDS) (\%)} + 0.4\%$ and used in the present study.

3.2.3. Study quality and risk of bias assessment

The Jadad Scale was adopted to assess the quality of all included studies. We evaluated the included studies regarding randomization, concealment of allocation, double-blinding, withdrawals and dropouts. To assess publication bias, we used the funnel plot and Egger’s test.

3.2.4. Statistical analysis

WMD and 95% confidence interval for HbA1c change from baseline in SGLT-2 inhibitors compared with placebo was calculated. Heterogeneity between studies was assessed by using Q statistic and I^2 statistic. A P -value of Q statistic <0.10 and a value of 50% or greater of I^2 statistic was considered significant. A random-effect model was used because significant heterogeneity was reported in several meta-analyses [26, 27]. Meta-regression was applied to investigate potential factors affecting the placebo adjusted HbA1c lowering efficacy. The pre-specified factors such as study duration, percentage of male, age, duration of diabetes, baseline HbA1c values, fasting plasma glucose, BMI, body weight, percentage of Asians and additional categorizations regarding ethnicity of study subjects and study regions were used in the meta-regression.

Univariate meta-regression was performed first to identify potential factors affecting the HbA1c-lowering efficacy of the SGLT-2 inhibitor. Negative coefficient indicates efficacy is in favor of the factor. A significant association was defined as $P <0.2$ in univariate meta-regression. Factors were further analyzed by multivariate meta-regression. When correlation between explanatory variables was identified, one variable was selected to be included in the multivariate meta-regression. A statistically significant difference was defined as $P <0.05$ in the multivariate meta-regression. Analyses were performed using StatDirect version 2.7.9 (StatDirect Ltd., Altrincham, Cheshire, UK) and R software, version 3.2.0 [23].

3.3. Result

3.3.1. Literature search

We identified 496 potentially relevant articles from a literature search. Based on the review of the abstracts, 29 studies met all of the inclusion criteria. Further 83 studies were identified in the ClinicalTrials.gov registry and among them 1 unpublished clinical study was included in the analysis (Fig. 3). A total of 30 studies (Appendix 2: 1-33) with 75 arms and 20,170 patients were used for the analysis (Table 4). Regarding the factors related to ethnicity and regions, in Type B, 6 were studies in Asian subjects, 5 were studies in Japanese subjects and 20 were studies in non-Asian subjects. As for Type C, 2 were studies conducted in Asia, 5 were studies conducted in Japan and 24 were studies conducted multinationally (Table 5).

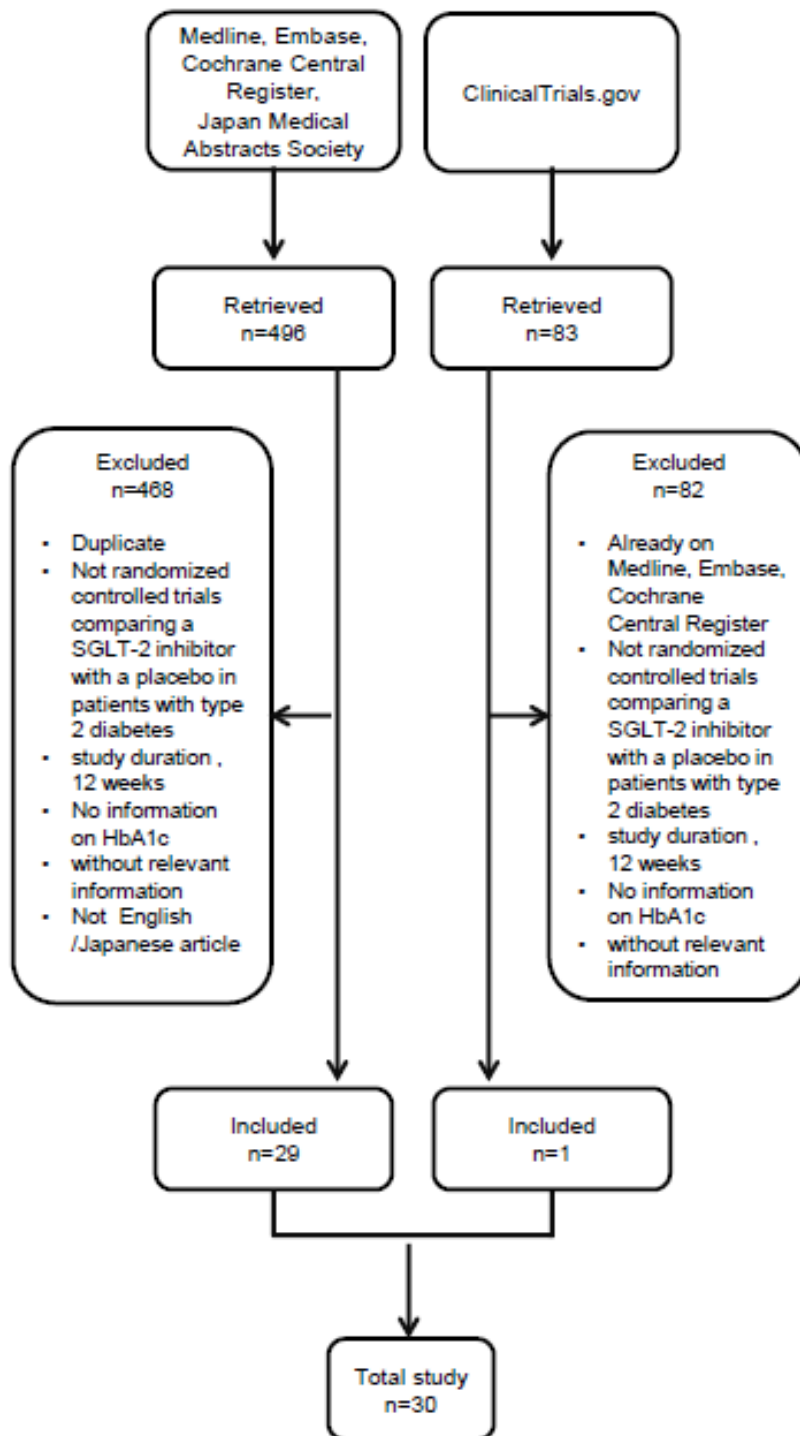


Figure 3 Flow diagram of study selection

Table 4 Studies included in the analysis

Source	Number of Treatment participants	duration (weeks)	SGLT-2 inhibitors	background antidiabetic therapy	Age (years)	Baseline HbA1c (%)	BL FPG (mg/dl)
List, 2009 ¹	112	12	Dapagliflozin 5 mg QD	none	54.0	8.0	151.6
List, 2009 ¹	101	12	Dapagliflozin 10 mg QD	none	53.5	8.0	149.1
Bailey, 2010 ²	274	24	Dapagliflozin 5 mg QD	metformin	54.0	8.1	167.2
Bailey, 2010 ²	272	24	Dapagliflozin 10 mg QD	metformin	53.2	8.0	160.7
Ferrannini, 2010 ³ _1 ^a	139	24	Dapagliflozin 5 mg QD	none	52.7	7.9	161.0
Ferrannini, 2010 ³ _1 ^a	145	24	Dapagliflozin 10 mg QD	none	51.7	7.9	163.1
Ferrannini, 2010 ³ _2 ^b	135	24	Dapagliflozin 5 mg QD	none	53.6	7.8	158.4
Ferrannini, 2010 ³ _2 ^b	143	24	Dapagliflozin 10 mg QD	none	51.6	7.9	164.3
Strojek, 2011 ⁴	287	24	Dapagliflozin 5 mg QD	glimepiride	60.3	8.1	173.3
Strojek, 2011 ⁴	296	24	Dapagliflozin 10 mg QD	glimepiride	59.6	8.1	172.2
Biley, 2012 ⁵	136	24	Dapagliflozin 5 mg QD	none	52.4	7.9	159.2
Rosenstock, 2012 ⁶	280	24	Dapagliflozin 5 mg QD	pioglitazone	53.3	8.4	164.7
Rosenstock, 2012 ⁶	279	24	Dapagliflozin 10 mg QD	pioglitazone	53.7	8.4	162.8
Kaku, 2013 ⁷	112	12	Dapagliflozin 5 mg QD	none	58.2	8.1	161.8
Kaku, 2013 ⁷	106	12	Dapagliflozin 10 mg QD	none	57.5	8.2	161.1

Source	Number of Treatment		SGLT-2 inhibitors	background antidiabetic therapy	Age (years)	Baseline HbA1c (%)	BL FPG (mg/dl)
	participants	(weeks)					
Jabbour, 2014 ⁸	221	24	Dapagliflozin 10 mg QD	sitagliptin	53.0	8.1	159.4
Jabbour, 2014 ⁸	226	24	Dapagliflozin 10 mg QD	sitagliptin and metformin	56.7	7.9	165.8
Kaku, 2014 ⁹	173	24	Dapagliflozin 5 mg QD	none	59.5	7.5	138.6
Kaku, 2014 ⁹	175	24	Dapagliflozin 10 mg QD	none	58.9	7.5	139.2
Ji, 2014 ¹⁰	260	24	Dapagliflozin 5 mg QD	none	51.4	8.3	160.8
Ji, 2014 ¹⁰	265	24	Dapagliflozin 10 mg QD	none	50.6	8.3	164.7
Matthaei, 2015 ¹¹	218	24	Dapagliflozin 10 mg QD	sulfonylurea and metformin	61.0	8.2	173.8
Schumm-Draeger, 2015 ¹²	201	16	Dapagliflozin 2.5 mg BID	metformin	58.4	7.9	155.4
Schumm-Draeger, 2015 ¹²	200	16	Dapagliflozin 5 mg BID	metformin	56.9	7.9	156.4
NCT00736879 ¹³	136	24	Dapagliflozin 5 mg QD	none	52.4	7.9	159.2
Rosenstock, 2012 ¹⁴	129	12	Canagliflozin 100 mg QD	metformin	52.5	7.8	166.0
Rosenstock, 2012 ¹⁴	130	12	Canagliflozin 200 mg QD	metformin	53.1	7.7	162.0
Rosenstock, 2012 ¹⁴	129	12	Canagliflozin 300 mg QD	metformin	52.8	7.7	161.5
Wilding, 2013 ¹⁵	313	26	Canagliflozin 100 mg QD	sulfonylurea and metformin	57.1	8.1	171.0
Wilding, 2013 ¹⁵	312	26	Canagliflozin 300 mg QD	sulfonylurea and metformin	56.5	8.1	168.3
Inagaki, 2013 ¹⁶	149	12	Canagliflozin 100 mg QD	none	57.7	8.0	165.9
Inagaki, 2013 ¹⁶	151	12	Canagliflozin 200 mg QD	none	57.3	8.1	168.3
Inagaki, 2013 ¹⁶	150	12	Canagliflozin 300 mg QD	none	57.4	8.1	169.9

Source	Number of Treatment participants	duration (weeks)	SGLT-2 inhibitors background antidiabetic therapy	Age (years)	Baseline HbA1c (%)	BL FPG (mg/dl)
Lavalle-Gonzalez, 2013 ¹⁷	551	26	Canagliflozin 100 mg QD metformin	55.4	7.9	166.2
Lavalle-Gonzalez, 2013 ¹⁷	550	26	Canagliflozin 300 mg QD metformin	55.3	7.9	169.8
Stenlof, 2013 ¹⁸	387	26	Canagliflozin 100 mg QD none	55.4	8.1	170.1
Stenlof, 2013 ¹⁸	389	26	Canagliflozin 300 mg QD none	55.5	8.0	170.1
Forst, 2014 ¹⁹	228	26	Canagliflozin 100 mg QD pioglitazone and metformin	57.5	8.0	166.5
Forst, 2014 ¹⁹	229	26	Canagliflozin 300 mg QD pioglitazone and metformin	57.7	8.0	163.8
Inagaki, 2014 ²⁰	183	24	Canagliflozin 100 mg QD none	58.3	8.0	160.4
Inagaki, 2014 ²⁰	182	24	Canagliflozin 200 mg QD none	57.8	8.0	164.1
Qiu, 2014 ²¹	186	18	Canagliflozin 50 mg BID metformin	57.8	7.7	162.0
Qiu, 2014 ²¹	186	18	Canagliflozin 150 mg metformin BID	56.9	7.7	162.9
Ji, 2015 ²² _1 ^a	449	18	Canagliflozin 100 mg QD metformin (with or without sulfonylurea)	56.1	7.9	157.5
Ji, 2015 ²² _1 ^a	453	18	Canagliflozin 300 mg QD metformin (with or without sulfonylurea)	56.1	8.0	159.3
Ji, 2015 ²² _2 ^b	218	18	Canagliflozin 100 mg QD metformin	54.7	7.9	158.4
Ji, 2015 ²² _2 ^b	222	18	Canagliflozin 300 mg QD metformin	55.3	8.0	158.4
Ji, 2015 ²² _3 ^c	231	18	Canagliflozin 100 mg QD sulfonylurea and metformin	57.6	7.9	155.7

Source	Number of Treatment		SGLT-2 inhibitors	background antidiabetic therapy	Age (years)	Baseline HbA1c (%)	BL FPG (mg/dl)
	participants	duration (weeks)					
Ji, 2015 ²² _3 ^c	231	18	Canagliflozin 300 mg QD	sulfonylurea and metformin	57.0	8.0	160.2
Haring, 2013 ²³	450	24	Empagliflozin 10 mg QD	sulfonylurea and metformin	57.0	8.1	151.2
Haring, 2013 ²³	441	24	Empagliflozin 25 mg QD	sulfonylurea and metformin	57.1	8.1	153.9
Roden, 2013 ²⁴	452	24	Empagliflozin 10 mg QD	none	55.5	7.9	153.6
Roden, 2013 ²⁴	452	24	Empagliflozin 25 mg QD	none	54.4	7.9	153.5
Rosenstock, 2013 ²⁵	142	12	Empagliflozin 10 mg QD	metformin	59.5	8.0	173.5
Rosenstock, 2013 ²⁵	141	12	Empagliflozin 25 mg QD	metformin	59.5	8.0	177.0
Ferrannini, 2013 ²⁶	163	12	Empagliflozin 10 mg QD	none	57.8	7.9	174.6
Ferrannini, 2013 ²⁶	164	12	Empagliflozin 25 mg QD	none	57.0	7.8	171.0
Kovacs, 2014 ²⁷	330	24	Empagliflozin 10 mg QD	pioglitazone (with or without metformin)	54.7	8.2	151.8
Kovacs, 2014 ²⁷	333	24	Empagliflozin 25 mg QD	pioglitazone (with or without metformin)	54.4	8.2	151.7
Haring, 2014 ²⁸	424	24	Empagliflozin 10 mg QD	metformin	55.7	7.9	155.1
Haring, 2014 ²⁸	421	24	Empagliflozin 25 mg QD	metformin	55.8	7.9	152.5
Kadowaki, 2014 ²⁹	218	12	Empagliflozin 10 mg QD	none	58.3	7.9	156.6
Kadowaki, 2014 ²⁹	218	12	Empagliflozin 25 mg QD	none	58.0	8.0	156.2
Ross, 2015 ³⁰	322	16	Empagliflozin 12.5 mg	metformin BID	57.7	7.8	157.8
Ross, 2015 ³⁰	321	16	Empagliflozin 25 mg QD	metformin	58.1	7.7	159.0

Source	Number of Treatment participants	duration (weeks)	SGLT-2 inhibitors	background antidiabetic therapy	Age (years)	Baseline HbA1c (%)	BL FPG (mg/dl)
Ross, 2015 ³⁰	322	16	Empagliflozin 5 mg BID	metformin	58.5	7.8	161.4
Ross, 2015 ³⁰	321	16	Empagliflozin 10 mg QD	metformin	58.3	7.8	161.4
Roden, 2015 ³¹	452	76	Empagliflozin 10 mg QD	none	55.5	7.9	153.9
Roden, 2015 ³¹	452	76	Empagliflozin 25 mg QD	none	54.4	7.9	153.9
Haering, 2015 ³²	450	76	Empagliflozin 10 mg QD	sulfonylurea and metformin	57.0	8.1	151.2
Haering, 2015 ³²	441	76	Empagliflozin 25 mg QD	sulfonylurea and metformin	57.1	8.1	153.9
NCT01289990 ³³ _1 ^a	424	76	Empagliflozin 10 mg QD	metformin	55.7	7.9	155.1
NCT01289990 ³³ _1 ^a	421	76	Empagliflozin 25 mg QD	metformin	55.8	7.9	152.5
NCT01289990 ³³ _2 ^b	331	76	Empagliflozin 10 mg QD	pioglitazon	55.0	NR	NR
NCT01289990 ³³ _2 ^b	334	76	Empagliflozin 25 mg QD	pioglitazon	55.1	NR	NR

QD once-daily dosing, *BID* twice-daily dosing, *NR* not reported

^a 1 represents the first pair of the article

^b 2 represents the second pair of the article

^c 3 represents the third pair of the article

Table 5 Three classifications for studies

Type A	Study in non-Asian subjects 19 studies	Study in Asian subjects 11 studies	
Type B	Study in non-Asian subjects 20 studies	Study in Asian subjects (excluding the Japanese) 6 studies	Study in Japanese subjects ^a 5 studies
Type C	Study conducted multinationally 24 studies	Study conducted in Asia (excluding Japan) 2 studies	Study conducted in Japan ^a 5 studies

^a Studies classified as "Study in Japanese subjects" (Type B) and "Study conducted in Japan" (Type C) were the same.

3.3.2. Quality of included studies and publication bias

The Jaded score was 3.90, and we determined the quality of included studies was high. The funnel plot of WMD for HbA1c and standard error in 30 studies is shown in Fig. 4. Funnel plot asymmetry was assessed using Egger's test and it indicated that there was no publication bias ($p = 0.117$). The I^2 statistic test for heterogeneity was 76.2%, and the value of Q statistic was significant.

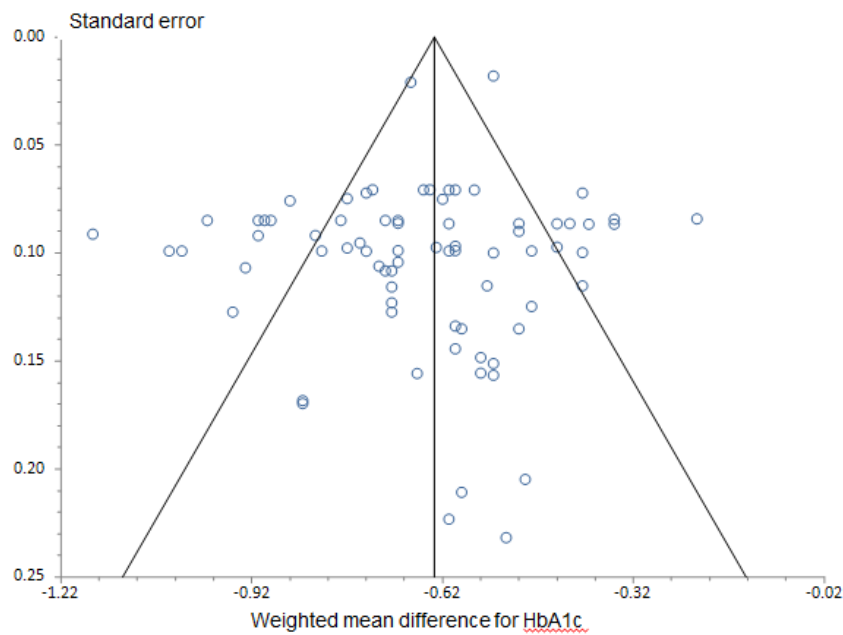


Figure 4 Funnel plot of weighted mean difference for HbA1c and standard error

3.3.3. Primary outcome

HbA1c data for 75 arms from 30 studies were pooled. The difference of HbA1c between the SGLT-2 inhibitor group and placebo group was -0.657% (95% CI -0.696 to -0.617).

3.3.4. Univariate meta-regression

Univariate meta-regression showed relations between the WMD in the change of HbA1c from baseline value and several factors such as study design (coefficient=0.201, $p < 0.001$), percentage of male (-0.009 , $p < 0.001$), baseline HbA1c values (-0.087 , $p = 0.162$), baseline fasting plasma glucose (-0.006 , $p = 0.020$), BMI (0.015 , $p = 0.081$), body weight (0.003 , $p = 0.180$) and percentage of Asians (-0.001 , $p = 0.011$). Additionally, studies in Asian subjects in Type A (-0.103 , $p = 0.016$), studies in Japanese subjects in Type B (-0.200 , $p = 0.001$), and studies conducted in Japan in Type C (-0.178 , $p = 0.002$) showed relations in each of the analyses. Because BMI and body weight correlated with the percentage of Asians, we excluded BMI and body weight from the multivariate meta-regression, and used the three classifications (Type A, B and C). Furthermore, study design correlated with the three classifications, we exclude study design from multivariate meta-regression.

3.3.5. Multivariate meta-regression

The factors significantly associated with the WMD in the change of HbA1c from baseline using univariate meta-regression were further analyzed by multivariate meta-regression (Table 6). Baseline fasting plasma glucose was identified as a common influencing factor in all the analyses using the three types of classifications. Also, percentage of male in Type A, and percentage of male and baseline HbA1c were identified as influencing factors, respectively.

Table 6 Results of multivariate meta-regression analysis

Factors	Estimate	95% CI	P-value
a. Multivariate meta-regression with Type A			
Male (%)	-0.007	(-0.012, -0.002)	0.008
Baseline HbA1c (%)	-0.111	(-0.357, 0.136)	0.378
Baseline fasting plasma glucose (mg/dl)	-0.009	(-0.015, -0.003)	0.003
Study in Asian subjects	-0.097	(-0.200, 0.006)	0.066
b. Multivariate meta-regression with Type B			
Male (%)	-0.005	(-0.011, 0.002)	0.142
Baseline HbA1c (%)	-0.138	(-0.394, 0.118)	0.292
Baseline fasting plasma glucose (mg/dl)	-0.009	(-0.015, -0.003)	0.005
Study in Asian subjects	-0.086	(-0.193, 0.021)	0.114
Study in Japanese subjects	-0.149	(-0.315, 0.016)	0.077
c. Multivariate meta-regression with Type C			
Male (%)	-0.007	(-0.014, 0.000)	0.036
Baseline HbA1c (%)	-0.237	(-0.470, -0.004)	0.046
Baseline fasting plasma glucose (mg/dl)	-0.006	(-0.011, -0.001)	0.029
Study conducted in Asia	0.070	(-0.044, 0.185)	0.227
Study conducted in Japan	-0.063	(-0.222, 0.096)	0.435

Type A, the studies were divided into 2 groups; a “study in Asian subjects” or a “study in non-Asian subjects”

Type B, the studies were divided into 3 groups; a “study in Asian subjects”, a “study in Japanese subjects” or a “study in non-Asian subjects.”

Type C, the studies were divided into three groups: a “study conducted in Asia”, a “study conducted in Japan” or a “study conducted multinationally.”

HbA1c hemoglobin A1c

3.4. Discussion

We performed meta-regression analyses to identify the factors influencing the glucose-lowering efficacy of SGLT-2 inhibitors by analyzing 30 studies with 75 arms

and 20,170 patients. We found WMD in the placebo-adjusted change of HbA1c from baseline was -0.66% , which is consistent with the previously reported meta-analyses ranging between -0.455% and -0.78% .

Using univariate meta-regression, factors such as percentage of male, baseline HbA1c values, baseline fasting plasma glucose, BMI, body weight and percentage of Asians were identified to be correlated with WMD in the change of HbA1c from baseline. The results of baseline HbA1c values, baseline fasting plasma glucose and BMI were in agreement with the analysis conducted by Monami et al. [27]. Additionally, in the present study, relations were shown in studies in Asian subjects in the classification Type A, studies in Japanese subjects in Type B and studies conducted in Japan in Type C. Based on these results, SGLT-2 inhibitors exhibit a better glucose-lowering efficacy with higher baseline HbA1c, higher fasting glucose, with lower BMI and body weight, and with a higher percentage of Asians. A better glucose lowering efficacy was also shown in studies in Asian subjects, studies in Japanese subjects and studies conducted in Japan.

Multivariate meta-regression showed a better glucose lowering efficacy with high baseline fasting plasma glucose, suggesting baseline fasting plasma glucose is a robust factor. Additional categorizations regarding ethnicity of study subjects and study regions were not identified as an influencing factor of glucose-lowering efficacy of SGLT-2 inhibitors.

Ferrannini et al. [28] concluded SGLT-2 inhibitor induced glycosuria could be described as a dual function of the glomerular filtration rate (GFR) and plasma glucose levels with a coefficient of determination of 65% (or a multiple $r = 0.74$). This may explain our findings that higher baseline fasting plasma glucose is a factor of greater

efficacy of SGLT-2 inhibitors on WMD in the change of HbA1c from baseline. Because GFR was assessed in a few clinical studies, we did not include GFR as a pre-specified factor. Our finding suggests that paying attention to baseline fasting plasma glucose in planning and conducting of clinical studies, and in comparing the data with other studies is important.

We used three classifications regarding ethnicity of study subjects and study regions to explore potential differences among them. In the classification Type A, Japan was included in Asia. In Type B and C, we distinguished between Asia and Japan focusing on ethnicity in Type B and on regions in Type C. Any factors of these classifications were not identified as an influencing factor of glucose-lowering efficacy of SGLT-2 inhibitors. Our finding indicates that differences in the contribution of the insulin secretory defect and insulin resistance in the pathophysiology of type 2 diabetes between Asians and non-Asians do not affect the response to SGLT-2 inhibitors. The reason could be explained by its insulin independent mechanism of action, neither insulin secretion nor sensitization.

Our analysis indicates that the baseline fasting plasma glucose is associated with the efficacy of SGLT-2 inhibitors. This result should be taken into account when planning and conducting clinical studies, and comparing the data with other clinical studies.

4. Research 3 (Placebo effect)

4.1. Objectives

Through our research 1 and 2, different placebo phenomenon was observed among study regions. For example, high placebo responses were observed in study conducted in Asia and low placebo responses were observed in study conducted in Japan. This placebo phenomenon showed same trend in both the DPP -4 inhibitor and the SGLT-2 inhibitor studies.

The placebo phenomenon in chronic pain or neuropsychiatric disorder has been widely recognized in clinical research because high placebo responses were reported to be the cause of failure of clinical studies. In type 2 diabetes area, however, placebo effect has rarely been discussed.

Thus, we conducted meta-regression analyses to assess the factors that contributing to placebo effect in clinical studies for both DPP-4 inhibitor and SGLT-2 inhibitor. Potential factors including ethnicity of study subjects and study regions; in particular, Japanese were dealt separately from other Asians.

4.2. Method

4.2.1. Data sources, searches and extraction

We used the eligible studies identified in research 1 and research 2. The primary outcome was HbA1c change from baseline in placebo group at each study's primary endpoint.

In order to further explore potential differences in HbA1c change from baseline in placebo group among the ethnicity and study regions, we used two classifications, Type B and C, described in method of research 1 and research 2.

4.2.2. Statistical analysis

Meta-regression was applied to investigate potential factors affecting the HbA1c change from baseline in placebo group of each DPP-4 inhibitor and SGLT-2 inhibitor study. The pre-specified factors such as study duration, percentage of male, age, duration of diabetes, baseline HbA1c values, fasting plasma glucose, BMI, body weight, percentage of Asians and additional categorizations regarding ethnicity of study subjects and study regions were used in the meta-regression.

Same method was applied for univariate and multivariate meta-regression analyses as in research 1 and 2. Analyses were performed using StatDirect version 2.7.9 (StatDirect Ltd., Altrincham, Cheshire, UK) and R software, version 3.2.0 [23].

4.3. Result

4.3.1. Literature search

DPP-4 inhibitor study

A total of 79 studies (Appendix 1; reference 1-79) with 86 arms and 10,561 patients were used for the analysis (Table 1). Regarding the factors related to ethnicity and regions, in Type B, 17 were studies in Asian subjects, 18 were studies in Japanese subjects and 41 were studies in non-Asian subjects. As for Type C, 6 were studies conducted in Asia, 18 were studies conducted in Japan and 55 were studies conducted multinationally (Table 2).

SGLT-2 inhibitor study

A total of 30 studies (Appendix 2: 1-33) with 38 arms and 4,904 patients were

used for the analysis (Table 4). Regarding the factors related to ethnicity and regions, in Type B, 6 were studies in Asian subjects, 5 were studies in Japanese subjects and 20 were studies in non-Asian subjects. As for Type C, 2 were studies conducted in Asia, 5 were studies conducted in Japan and 24 were studies conducted multinationally (Table 5).

4.3.2. Univariate meta-regression

DPP-4 inhibitor study

Univariate meta-regression showed relations between the HbA1c change from baseline in placebo group and several factors such as study design (coefficient=-0.112, $p = 0.030$), study duration (-0.007 , $p = 0.052$), percentage of male (0.007 , $p = 0.007$), age (0.031 , $p < 0.001$), duration of diabetes (0.031 , $p = 0.015$), BMI (-0.012 , $p = 0.174$) and body weight (-0.005 , $p = 0.173$). Additionally, studies in Japanese subjects (0.198 , $p < 0.001$) and studies in Asian subjects (-0.173 , $p = 0.002$) in Type B, and studies conducted in Japan (0.228 , $p < 0.001$) and studies conducted in Asia (-0.259 , $p = 0.001$) in Type C showed relations in each of the analyses. Because BMI and body weight correlated with the percentage of Asians, we excluded BMI and body weight from the multivariate meta-regression, and used the two classifications (Type B and C).

SGLT-2 inhibitor study

Univariate meta-regression showed relations between the HbA1c change from baseline in placebo group and several factors such as study design (coefficient=-0.119, $p = 0.001$), percentage of male (0.008 , $p = 0.024$), age (0.025 , $p = 0.051$) and percentage of Asians (0.001 , $p = 0.163$). Additionally, studies in Japanese subjects in Type B (0.316 ,

$p < 0.001$), and studies conducted in Japan (0.292, $p < 0.001$) and studies conducted in Asia (-0.299 , $p = 0.001$) in Type C showed relations in each of the analyses.

4.3.3. Multivariate meta-regression

DPP-4 inhibitor study

The factors significantly associated with the HbA1c change from baseline in placebo group using univariate meta-regression were further analyzed by multivariate meta-regression (Table 7). Study design and duration of diabetes were identified as a common influencing factor in the both analyses using the two types of classifications. Also, study in Asian subjects in Type B and study conducted in Asia in Type C were identified as influencing factors.

Table 7 Results of multivariate meta-regression analysis (*DPP-4 inhibitor study*)

Factors	Estimate	95% CI	P-value
a. Multivariate meta-regression with Type B			
Study design (mono/combination)	-0.192	(-0.356, -0.027)	0.022
Study duration (week)	-0.004	(-0.018, 0.010)	0.571
Male (%)	0.001	(-0.007, 0.009)	0.820
Age (years)	-0.005	(-0.028, 0.018)	0.643
Duration of diabetes (years)	0.046	(0.003, 0.089)	0.036
Study in Asian subjects	-0.185	(-0.336, -0.034)	0.017
Study in Japanese subjects	0.104	(-0.078, 0.286)	0.261
b. Multivariate meta-regression with Type C			
Study design (mono/combination)	-0.202	(-0.362, -0.043)	0.013
Study duration (week)	-0.004	(-0.018, 0.009)	0.542
Male (%)	0.001	(-0.007, 0.009)	0.834
Age (years)	-0.002	(-0.025, 0.020)	0.845
Duration of diabetes (years)	0.049	(0.007, 0.091)	0.021
Study conducted in Asia	-0.235	(-0.414, -0.056)	0.010
Study conducted in Japan	0.099	(-0.082, -0.279)	0.284

Type B, the studies were divided into 3 groups; a “study in Asian subjects”, a “study in Japanese subjects” or a “study in non-Asian subjects.”

Type C, the studies were divided into three groups: a “study conducted in Asia”, a “study conducted in Japan” or a “study conducted multinationally.”

HbA1c hemoglobin A1c

SGLT-2 inhibitor study

The factors significantly associated with the HbA1c change from baseline in placebo group using univariate meta-regression were further analyzed by multivariate meta-regression (Table 8). Study design was identified as a common influencing factor in the both analyses using the two types of classifications. Also, age with type B and study conducted in Asia in Type C was identified as influencing factors.

Table 8 Results of multivariate meta-regression analysis (*SGLT-2 inhibitor study*)

Factors	Estimate	95% CI	P-value
a. Multivariate meta-regression with Type B			
Study design (mono/combination)	-0.236	(-0.380, -0.092)	0.001
Male (%)	0.000	(-0.009, 0.008)	0.923
Age	0.035	(0.007, 0.062)	0.013
Study in Asian subjects	0.000	(-0.128, 0.128)	0.998
Study in Japanese subjects	0.083	(-0.192, 0.358)	0.553
b. Multivariate meta-regression with Type C			
Study design (mono/combination)	-0.211	(-0.339, -0.083)	0.001
Male (%)	0.003	(-0.005, 0.012)	0.435
Age	0.025	(0.000, 0.050)	0.052
Study conducted in Asia	-0.262	(-0.048, -0.076)	0.006
Study conducted in Japan	0.032	(-0.213, 0.276)	0.798

Type B, the studies were divided into 3 groups; a “study in Asian subjects”, a “study in Japanese subjects” or a “study in non-Asian subjects.”

Type C, the studies were divided into three groups: a “study conducted in Asia”, a “study conducted in Japan” or a “study conducted multinationally.”

HbA1c hemoglobin A1c

4.4. Discussion

In the present study, we conducted analysis of placebo effect in clinical studies for DPP-4 inhibitors and SGLT-2 inhibitors. Multivariate meta-regression showed a larger HbA1c change from baseline in placebo group in the studies conducted in Asia for the both analyses, suggesting studies conducted in Asia is a robust factor. In other words, larger placebo response, improvement in the placebo group, was shown in studies conducted in Asia. This result was in agreement with what we found in our research 1 and 2.

He et al. [29] conducted meta-analysis and reported that HbA1c in the placebo arm declined by 0.26 % in the studies of DPP-4 inhibitors conducted in patients with

type 2 diabetes in China, whereas the placebo effect of those conducted outside China was close to 0. The authors concluded that there were significant differences in response in the placebo group of DPP-4 studies conducted in China and those conducted outside of China. In our research, China was the main region in the studies in Asia. Our findings were in agreement with those by He et al. [29], not only for DPP-4 studies but also for SGLT-2 studies, and supported their consideration of the reason for a large placebo effect; the significant benefit to participate in clinical study for subjects in China, obtaining more resource of medical care, could cause a bias and obvious impact on management of diabetes, and thus cause better blood glucose control even in placebo arm in China.

Although there was no statistical difference, coefficients of studies in Japanese subjects in Type B and studies conducted in Japan in Type C were positive and had the least influence on placebo effect. This is thought to be due to the adequate management in the medical practice in Japan, which causes less impact on placebo arm.

5. Overall Discussion

In order to discuss the points to be considered in future clinical development of oral hypoglycemic drug, three types of research were conducted.

The finding of the first research suggested that a higher baseline HbA1c is a factor of greater efficacy of DPP-4 inhibitors on WMD in the change of HbA1c from baseline. A higher baseline HbA1c has been reported to be a predictor of a greater HbA1c reduction with the use of DPP-4 inhibitors [20, 25]. However, the HbA1c value used by Esposito et al. [20] and Deacon [25] in their analysis was different from our study; they used absolute HbA1c reduction without taking into account the placebo effect. We used the WMD in the change of HbA1c in multivariate meta-regression and the analysis showed a better glucose lowering efficacy with high baseline HbA1c, suggesting baseline HbA1c is a robust factor.

Second, studies in Japanese subjects and studies conducted in Japan are factors affecting the efficacy of DPP-4 inhibitors. However, studies in Asian subjects and studies conducted in Asia are not associated, which means the efficacy of DPP-4 inhibitors shows no apparent difference between studies in Asian subjects and studies in non-Asian subject, or studies conducted in Asian and studies conducted multinationally. This finding differs from those by Kim et al. [15], who reported that the DPP-4 inhibitors exhibit a better glucose-lowering efficacy in Asians than in other ethnic groups. Our finding indicates that differences in the contribution of the insulin secretory defect and the insulin resistance in the pathophysiology of type 2 diabetes between Asians and non-Asians do not significantly affect the response to DPP-4 inhibitors and it does not explain the mechanism underlying the difference in the response.

The finding of the second research suggested that a higher baseline fasting

plasma glucose is a factor of greater efficacy of SGLT-2 inhibitors on WMD in the change of HbA1c from baseline. It was based on the multivariate meta-regression, suggesting baseline fasting plasma glucose is a robust factor. This finding is in agreement of the research by Monami et al. [27], though they employed a univariate meta-regression analysis. Our findings may be explained by the fact that glycosuria and plasma glucose levels have a strong positive correlation.

We used three classifications regarding ethnicity of study subjects and study regions to explore potential differences among them. However, no factors of these classifications were identified as an influencing factor of glucose-lowering efficacy of SGLT-2 inhibitors. Our finding indicates that differences in the contribution of the insulin secretory defect and the insulin resistance in the pathophysiology of type 2 diabetes between Asians and non-Asians does not affect the response to SGLT-2 inhibitors. The reason could be explain by its insulin independent mechanism of action, neither insulin secretion nor sensitization.

The finding of the third research suggested that studies conducted in Asia is a factor of larger HbA1c change from baseline in the placebo group. He et al. [29] reported that there were significant differences in response in the placebo group between DPP-4 studies conducted in China and those conducted outside of China. In our research, China was the main region in studies conducted in Asia. Given that our findings were in agreement with those by He et al. [29], not only for DPP-4 studies but also for SGLT-2 studies and supported their consideration of the reason for this high placebo effect; the significant benefit to participate in clinical study for subjects in China, obtaining more resource of medical care, could cause a bias and obvious impact on management of diabetes, and thus cause better blood glucose control even in placebo

arm in China.

Publication bias may exist and have imposed some limitations on our study. Another limitation is that we used summary data and not individual patient-level data. Furthermore, we included South Asians in the studies in Asian subjects, which mainly consisted of East Asians, because it was not possible to distinguish in some studies whether Asians included South Asians or not. The Asian population is also ethnically heterogeneous and has different demographic characteristics [30].

6. Conclusion

Paying attention to baseline HbA1c for DPP-4 inhibitors and baseline fasting plasma glucose for SGLT-2 inhibitors in planning and conducting clinical studies, and in comparing the data with other clinical studies is important. With regards to the ethnicity of study subjects, our finding indicates that differences in the contribution of the insulin secretory defect and the insulin resistance in the pathophysiology of type 2 diabetes between Asians and non-Asians do not affect the response to active oral hypoglycemic agents which have insulin independent mechanism. On the other hand, placebo effect was shown to be different among the subject ethnicity and/or study region. This differential HbA1c response in the placebo arm should be taken into consideration when comparing the data based on different clinical practice in each region. In the future, it is desirable to evaluate the influence of different study regions and/or subject ethnicity on efficacy and safety in multiregional studies.

References

1. World Health Organization Global report on diabetes. 2016. http://apps.who.int/iris/bitstream/10665/204871/1/9789241565257_eng.pdf?ua=1&ua=1. Accessed 14 June 2016
2. Guariguata L, Whiting DR, Hambleton I, Beagley J, Linnenkamp U, Shaw JE. Global estimates of diabetes prevalence for 2013 and projections for 2035. *Diabetes Res Clin Pract.* 2014;103(2):137-49. doi:10.1016/j.diabres.2013.11.002.
3. International Diabetes Federation Diabetes Atlas (Seventh edition). 2015. <http://www.diabetesatlas.org/component/attachments/?task=download&id=116>. Accessed 14 June 2016
4. Survey of the state of diabetes in Japan in 1997. 1997. [in Japanese] http://www.mhlw.go.jp/toukei/kouhyo/indexkk_4_1.html. Accessed 14-June 2016
5. National Health and Nutrition Survey Japan, 2012. 2012. [in Japanese] <http://www.mhlw.go.jp/bunya/kenkou/eiyou/h24-houkoku.html>. Accessed 14-June 2016
6. Kodama K, Tojjar D, Yamada S, Toda K, Patel CJ, Butte AJ. Ethnic Differences in the Relationship between Insulin Sensitivity and Insulin Response. *Diabetes Care.* 2013;36(6):1789-96.
7. DeFronzo RA. Pathogenesis of type 2 diabetes mellitus. *Med Clin North Am.* 2004;88(4):787-835. doi:http://dx.doi.org/10.1016/j.mcna.2004.04.013.
8. Kim DJ, Lee MS, Kim KW, Lee MK. Insulin secretory dysfunction and insulin resistance in the pathogenesis of Korean type 2 diabetes mellitus. *Metabolism.* 2001;50(5):590-3. doi:10.1053/meta.2001.22558.
9. Qian L, Xu L, Wang X, Fu X, Gu Y, Lin F et al. Early insulin secretion failure leads to diabetes in Chinese subjects with impaired glucose regulation. *Diabetes Metab Res Rev.* 2009;25(2):144-9. doi:10.1002/dmrr.922.
10. Matsumoto K, Miyake S, Yano M, Ueki Y, Yamaguchi Y, Akazawa S et al. Glucose tolerance, insulin secretion, and insulin sensitivity in nonobese and obese Japanese subjects. *Diabetes Care.* 1997;20(10):1562-8.
11. Fukushima M, Usami M, Ikeda M, Nakai Y, Taniguchi A, Matsuura T et al. Insulin secretion and insulin sensitivity at different stages of glucose tolerance: a cross-sectional study of Japanese type 2 diabetes. *Metabolism.* 2004;53(7):831-5.
12. Flatt PR, Bailey CJ, Green BD. Dipeptidyl peptidase IV (DPP IV) and related molecules in type 2 diabetes. *Frontiers in bioscience: a journal and virtual library.* 2008;13:3648-60.

13. Mulvihill EE, Drucker DJ. Pharmacology, physiology, and mechanisms of action of dipeptidyl peptidase-4 inhibitors. *Endocr Rev.* 2014;35(6):992-1019.
14. Seino Y, Kuwata H, Yabe D. Incretin-based drugs for type 2 diabetes: Focus on East Asian perspectives. *Journal of diabetes investigation.* 2016;7 Suppl 1:102-9.
15. Kim YG, Hahn S, Oh TJ, Kwak SH, Park KS, Cho YM. Differences in the glucose-lowering efficacy of dipeptidyl peptidase-4 inhibitors between Asians and non-Asians: a systematic review and meta-analysis. *Diabetologia.* 2013;56(4):696-708. doi:10.1007/s00125-012-2827-3.
16. Park H, Park C, Kim Y, Rascati KL. Efficacy and safety of dipeptidyl peptidase-4 inhibitors in type 2 diabetes: meta-analysis. *Ann Pharmacother.* 2012;46(11):1453-69. doi:10.1345/aph.1R041.
17. Davidson JA, Aguilar R, Lavallo Gonzalez FJ, Trujillo A, Alba M, Vijapurkar U et al. Efficacy and Safety of Canagliflozin in Type 2 Diabetes Patients of Different Ethnicity. *Ethnicity & disease.* 2016;26(2):221-8. doi:10.18865/ed.26.2.221.
18. Monami M, Cremasco F, Lamanna C, Marchionni N, Mannucci E. Predictors of response to dipeptidyl peptidase-4 inhibitors: evidence from randomized clinical trials. *Diabetes Metab Res Rev.* 2011;27(4):362-72. doi:10.1002/dmrr.1184.
19. Esposito K, Cozzolino D, Bellastella G, Maiorino MI, Chiodini P, Ceriello A et al. Dipeptidyl peptidase-4 inhibitors and HbA1c target of <7% in type 2 diabetes: meta-analysis of randomized controlled trials. *Diabetes Obes Metab.* 2011;13(7):594-603. doi:10.1111/j.1463-1326.2011.01380.x.
20. Esposito K, Chiodini P, Capuano A, Maiorino MI, Bellastella G, Giugliano D. Baseline glycemic parameters predict the hemoglobin A1c response to DPP-4 inhibitors: meta-regression analysis of 78 randomized controlled trials with 20,053 patients. *Endocrine.* 2014;46(1):43-51. doi:10.1007/s12020-013-0090-0.
21. Liberati A, Altman DG, Tetzlaff J, Mulrow C, Gotzsche PC, Ioannidis JP et al. The PRISMA statement for reporting systematic reviews and meta-analyses of studies that evaluate healthcare interventions: explanation and elaboration. *BMJ.* 2009;339:b2700. doi:10.1136/bmj.b2700.
22. Kashiwagi A, Kasuga M, Araki E, Oka Y, Hanafusa T, Ito H et al. International clinical harmonization of glycated hemoglobin in Japan: From Japan Diabetes Society to National Glycohemoglobin Standardization Program values. *Journal of Diabetes Investigation.* 2012;3(1):39-40.
23. R Core Team. R: A language and environment for statistical computing. R Foundation for Statistical Computing, Vienna, Austria. 2015. <http://www.R-project.org/>.

24. Bloomgarden ZT, Dodis R, Viscoli CM, Holmboe ES, Inzucchi SE. Lower baseline glycemia reduces apparent oral agent glucose-lowering efficacy: a meta-regression analysis. *Diabetes Care*. 2006;29(9):2137-9. doi:10.2337/dc06-1120.
25. Deacon CF. Dipeptidyl peptidase-4 inhibitors in the treatment of type 2 diabetes: a comparative review. *Diabetes, Obesity and Metabolism*. 2011;13(1):7-18. doi:10.1111/j.1463-1326.2010.01306.x.
26. Vasilakou D, Karagiannis T, Athanasiadou E, Mainou M, Liakos A, Bekiari E et al. Sodium-glucose cotransporter 2 inhibitors for type 2 diabetes: a systematic review and meta-analysis. *Annals of internal medicine*. 2013;159(4):262-74.
27. Monami M, Nardini C, Mannucci E. Efficacy and safety of sodium glucose co-transport-2 inhibitors in type 2 diabetes: a meta-analysis of randomized clinical trials - SGLT-2 inhibitors in type 2 diabetes. *Diabetes, Obesity and Metabolism*. 2014;16(5):457-66.
28. Ferrannini E, Veltkamp SA, Smulders RA, Kadokura T. Renal Glucose Handling. *Diabetes Care*. 2013;36(5):1260-5.
29. He L, Liu S, Shan C, Tu Y, Li Z, Zhang XD. Differential HbA1c response in the placebo arm of DPP-4 inhibitor clinical trials conducted in China compared to other countries: a systematic review and meta-analysis. *BMC pharmacology & toxicology*. 2016;17(1):40. doi:10.1186/s40360-016-0084-7.
30. Singh AK. Incretin response in Asian type 2 diabetes: Are Indians different? *Indian J Endocrinol Metab*. 2015;19(1):30-8. doi:10.4103/2230-8210.146861.

Acknowledgement

First and foremost I would like to express my sincere gratitude to Professor Mamoru Narukawa for his patience, motivation, immense knowledge, and thoughtful guidance of my research. His guidance helped me in all the time of research and writing this thesis. Without his guidance and consistent feedback, my research would not have been achievable.

I give special thanks to Mr. Masayuki Kaneko for instructing, supporting and encouraging me with valuable feedback and advice, especially from the statistical point of view. I am also grateful to Ms. Takako Nakata for kindly providing me with an administrative support. I would like to thank all colleagues in our department for a fresh incentive, warm encouragement and support through my Ph.D. course.

I am very grateful to Pfizer Japan for making my Ph.D. research possible, and I am especially obliged to Dr. Akihisa Harada, Dr. Marie-Pierre Hellio Le Graverand-Gastineau, Taro Ishibashi and Masahito Nagashima for the support and encouragement.

Finally but certainly not least, I am extremely grateful to my family, especially to my husband for continuous support and encouragement throughout my PhD course.

Appendix 1

1. DeFronzo RA, Fleck PR, Wilson CA, Mekki Q, Alogliptin Study G. Efficacy and safety of the dipeptidyl peptidase-4 inhibitor alogliptin in patients with type 2 diabetes and inadequate glycemic control: a randomized, double-blind, placebo-controlled study. *Diabetes Care*. 2008;31(12):2315-7. doi:10.2337/dc08-1035.
2. Nauck MA, Ellis GC, Fleck PR, Wilson CA, Mekki Q, Alogliptin Study G. Efficacy and safety of adding the dipeptidyl peptidase-4 inhibitor alogliptin to metformin therapy in patients with type 2 diabetes inadequately controlled with metformin monotherapy: a multicentre, randomised, double-blind, placebo-controlled study. *Int J Clin Pract*. 2009;63(1):46-55. doi:10.1111/j.1742-1241.2008.01933.x.
3. Pratley RE, Reusch JE, Fleck PR, Wilson CA, Mekki Q, Alogliptin Study G. Efficacy and safety of the dipeptidyl peptidase-4 inhibitor alogliptin added to pioglitazone in patients with type 2 diabetes: a randomized, double-blind, placebo-controlled study. *Curr Med Res Opin*. 2009;25(10):2361-71. doi:10.1185/03007990903156111.
4. Pratley RE, Kipnes MS, Fleck PR, Wilson C, Mekki Q, Alogliptin Study G. Efficacy and safety of the dipeptidyl peptidase-4 inhibitor alogliptin in patients with type 2 diabetes inadequately controlled by glyburide monotherapy. *Diabetes Obes Metab*. 2009;11(2):167-76. doi:10.1111/j.1463-1326.2008.01016.x.
5. Seino Y, Fujita T, Hiroi S, Hirayama M, Kaku K. Efficacy and safety of alogliptin in Japanese patients with type 2 diabetes mellitus: a randomized, double-blind, dose-ranging comparison with placebo, followed by a long-term extension study. *Curr Med Res Opin*. 2011;27(9):1781-92. doi:10.1185/03007995.2011.599371.
6. Kaku K, Itayasu T, Hiroi S, Hirayama M, Seino Y. Efficacy and safety of alogliptin added to pioglitazone in Japanese patients with type 2 diabetes: a randomized, double-blind, placebo-controlled trial with an open-label long-term extension study. *Diabetes Obes Metab*. 2011;13(11):1028-35. doi:10.1111/j.1463-1326.2011.01460.x.
7. Seino Y, Fujita T, Hiroi S, Hirayama M, Kaku K. Alogliptin plus voglibose in Japanese patients with type 2 diabetes: a randomized, double-blind, placebo-controlled trial with an open-label, long-term extension. *Curr Med Res Opin*. 2011;27 Suppl 3:21-9. doi:10.1185/03007995.2011.614936.
8. Seino Y, Hiroi S, Hirayama M, Kaku K. Efficacy and safety of alogliptin added to

- sulfonylurea in Japanese patients with type 2 diabetes: A randomized, double-blind, placebo-controlled trial with an open-label, long-term extension study. *J Diabetes Investig.* 2012;3(6):517-25. doi:10.1111/j.2040-1124.2012.00226.x.
9. Seino Y, Miyata Y, Hiroi S, Hirayama M, Kaku K. Efficacy and safety of alogliptin added to metformin in Japanese patients with type 2 diabetes: a randomized, double-blind, placebo-controlled trial with an open-label, long-term extension study. *Diabetes Obes Metab.* 2012;14(10):927-36. doi:10.1111/j.1463-1326.2012.01620.x.
 10. Pan C, Han P, Ji Q, Li C, Lu J, Yang J et al. Efficacy and safety of alogliptin in patients with type 2 diabetes mellitus: A multicentre randomized double-blind placebo-controlled Phase 3 study in mainland China, Taiwan, and Hong Kong. *J Diabetes.* 2016. doi:10.1111/1753-0407.12425.
 11. Forst T, Uhlig-Laske B, Ring A, Graefe-Mody U, Friedrich C, Herbach K et al. Linagliptin (BI 1356), a potent and selective DPP-4 inhibitor, is safe and efficacious in combination with metformin in patients with inadequately controlled Type 2 diabetes. *Diabet Med.* 2010;27(12):1409-19. doi:10.1111/j.1464-5491.2010.03131.x.
 12. Del Prato S, Barnett AH, Huisman H, Neubacher D, Woerle HJ, Dugi KA. Effect of linagliptin monotherapy on glycaemic control and markers of beta-cell function in patients with inadequately controlled type 2 diabetes: a randomized controlled trial. *Diabetes Obes Metab.* 2011;13(3):258-67. doi:10.1111/j.1463-1326.2010.01350.x.
 13. Taskinen MR, Rosenstock J, Tamminen I, Kubiak R, Patel S, Dugi KA et al. Safety and efficacy of linagliptin as add-on therapy to metformin in patients with type 2 diabetes: a randomized, double-blind, placebo-controlled study. *Diabetes Obes Metab.* 2011;13(1):65-74. doi:10.1111/j.1463-1326.2010.01326.x.
 14. Owens DR, Swallow R, Dugi KA, Woerle HJ. Efficacy and safety of linagliptin in persons with type 2 diabetes inadequately controlled by a combination of metformin and sulphonylurea: a 24-week randomized study. *Diabet Med.* 2011;28(11):1352-61. doi:10.1111/j.1464-5491.2011.03387.x.
 15. Lewin AJ, Arvay L, Liu D, Patel S, von Eynatten M, Woerle HJ. Efficacy and tolerability of linagliptin added to a sulfonylurea regimen in patients with inadequately controlled type 2 diabetes mellitus: an 18-week, multicenter, randomized, double-blind, placebo-controlled trial. *Clin Ther.* 2012;34(9):1909-19 e15. doi:10.1016/j.clinthera.2012.07.008.
 16. Kawamori R, Inagaki N, Araki E, Watada H, Hayashi N, Horie Y et al. Linagliptin monotherapy provides superior glycaemic control versus placebo or voglibose with comparable safety in Japanese patients with type 2 diabetes: a randomized, placebo

- and active comparator-controlled, double-blind study. *Diabetes Obes Metab.* 2012;14(4):348-57. doi:10.1111/j.1463-1326.2011.01545.x.
17. Ross SA, Rafeiro E, Meinicke T, Toorawa R, Weber-Born S, Woerle HJ. Efficacy and safety of linagliptin 2.5 mg twice daily versus 5 mg once daily in patients with type 2 diabetes inadequately controlled on metformin: a randomised, double-blind, placebo-controlled trial. *Curr Med Res Opin.* 2012;28(9):1465-74. doi:10.1185/03007995.2012.714360.
 18. Haak T, Meinicke T, Jones R, Weber S, von Eynatten M, Woerle HJ. Initial combination of linagliptin and metformin improves glycaemic control in type 2 diabetes: a randomized, double-blind, placebo-controlled study. *Diabetes Obes Metab.* 2012;14(6):565-74. doi:10.1111/j.1463-1326.2012.01590.x.
 19. Barnett AH, Patel S, Harper R, Toorawa R, Thiemann S, von Eynatten M et al. Linagliptin monotherapy in type 2 diabetes patients for whom metformin is inappropriate: an 18-week randomized, double-blind, placebo-controlled phase III trial with a 34-week active-controlled extension. *Diabetes Obes Metab.* 2012;14(12):1145-54. doi:10.1111/dom.12011.
 20. Bajaj M, Gilman R, Patel S, Kempthorne-Rawson J, Lewis-D'Agostino D, Woerle HJ. Linagliptin improved glycaemic control without weight gain or hypoglycaemia in patients with type 2 diabetes inadequately controlled by a combination of metformin and pioglitazone: a 24-week randomized, double-blind study. *Diabet Med.* 2014;31(12):1505-14. doi:10.1111/dme.12495.
 21. Chen Y, Ning G, Wang C, Gong Y, Patel S, Zhang C et al. Efficacy and safety of linagliptin monotherapy in Asian patients with inadequately controlled type 2 diabetes mellitus: A multinational, 24-week, randomized, clinical trial. *J Diabetes Investig.* 2015;6(6):692-8. doi:10.1111/jdi.12346.
 22. Wang W, Yang J, Yang G, Gong Y, Patel S, Zhang C et al. Efficacy and safety of linagliptin in Asian patients with type 2 diabetes mellitus inadequately controlled by metformin: A multinational 24-week, randomized clinical trial. *J Diabetes.* 2016;8(2):229-37. doi:10.1111/1753-0407.12284.
 23. J. Rosenstock J, Sankoh S, List JF. Glucose-lowering activity of the dipeptidyl peptidase-4 inhibitor saxagliptin in drug-naive patients with type 2 diabetes. *Diabetes Obes Metab.* 2008;10(5):376-86. doi:10.1111/j.1463-1326.2008.00876.x.
 24. Rosenstock J, Aguilar-Salinas C, Klein E, Nepal S, List J, Chen R et al. Effect of saxagliptin monotherapy in treatment-naive patients with type 2 diabetes. *Curr Med Res Opin.* 2009;25(10):2401-11. doi:10.1185/03007990903178735.
 25. DeFronzo RA, Hissa MN, Garber AJ, Luiz Gross J, Yuyan Duan R, Ravichandran S

- et al. The efficacy and safety of saxagliptin when added to metformin therapy in patients with inadequately controlled type 2 diabetes with metformin alone. *Diabetes Care*. 2009;32(9):1649-55. doi:10.2337/dc08-1984.
26. Hollander P, Li J, Allen E, Chen R, Investigators CV. Saxagliptin added to a thiazolidinedione improves glycemic control in patients with type 2 diabetes and inadequate control on thiazolidinedione alone. *J Clin Endocrinol Metab*. 2009;94(12):4810-9. doi:10.1210/jc.2009-0550.
 27. Chacra AR, Tan GH, Apanovitch A, Ravichandran S, List J, Chen R et al. Saxagliptin added to a submaximal dose of sulphonylurea improves glycaemic control compared with uptitration of sulphonylurea in patients with type 2 diabetes: a randomised controlled trial. *Int J Clin Pract*. 2009;63(9):1395-406. doi:10.1111/j.1742-1241.2009.02143.x.
 28. Yang W, Pan CY, Tou C, Zhao J, Gause-Nilsson I. Efficacy and safety of saxagliptin added to metformin in Asian people with type 2 diabetes mellitus: a randomized controlled trial. *Diabetes Res Clin Pract*. 2011;94(2):217-24. doi:10.1016/j.diabres.2011.07.035.
 29. Pan CY, Yang W, Tou C, Gause-Nilsson I, Zhao J. Efficacy and safety of saxagliptin in drug-naive Asian patients with type 2 diabetes mellitus: a randomized controlled trial. *Diabetes Metab Res Rev*. 2012;28(3):268-75. doi:10.1002/dmrr.1306.
 30. Frederich R, McNeill R, Berglind N, Fleming D, Chen R. The efficacy and safety of the dipeptidyl peptidase-4 inhibitor saxagliptin in treatment-naive patients with type 2 diabetes mellitus: a randomized controlled trial. *Diabetol Metab Syndr*. 2012;4(1):36. doi:10.1186/1758-5996-4-36.
 31. White JL, Buchanan P, Li J, Frederich R. A randomized controlled trial of the efficacy and safety of twice-daily saxagliptin plus metformin combination therapy in patients with type 2 diabetes and inadequate glycemic control on metformin monotherapy. *BMC Endocr Disord*. 2014;14:17. doi:10.1186/1472-6823-14-17.
 32. Moses RG, Kalra S, Brook D, Sockler J, Monyak J, Visvanathan J et al. A randomized controlled trial of the efficacy and safety of saxagliptin as add-on therapy in patients with type 2 diabetes and inadequate glycaemic control on metformin plus a sulphonylurea. *Diabetes Obes Metab*. 2014;16(5):443-50. doi:10.1111/dom.12234.
 33. Seino Y, Efficacy and safety of saxagliptin in Japanese patients with type 2 diabetes - Two multi-centre, randomized, double-blind, placebo-controlled studies. *Japanese Pharmacology and Therapeutics* 2014;42(7):503-18. [in Japanese]
 34. Matthaei S, Catrinoiu D, Celiński A, Ekholm E, Cook W, Hirshberg B et al.

- Randomized, Double-Blind Trial of Triple Therapy With Saxagliptin Add-on to Dapagliflozin Plus Metformin in Patients With Type 2 Diabetes. *Diabetes Care*. 2015;38(11):2018-24. doi:10.2337/dc15-0811.
35. Evaluate Saxagliptin in Adult Indian Patients With Type 2 Diabetes Inadequate Glycemic Control.
<https://clinicaltrials.gov/ct2/show/NCT00918879?term=NCT00918879&rank=1>.
Accessed 7 July 2016
 36. Ristic S, Byiers S, Foley J, Holmes D. Improved glycaemic control with dipeptidyl peptidase-4 inhibition in patients with type 2 diabetes: vildagliptin (LAF237) dose response. *Diabetes Obes Metab*. 2005;7(6):692-8.
doi:10.1111/j.1463-1326.2005.00539.x.
 37. Dejager S, Razac S, Foley JE, Schweizer A. Vildagliptin in drug-naive patients with type 2 diabetes: a 24-week, double-blind, randomized, placebo-controlled, multiple-dose study. *Horm Metab Res*. 2007;39(3):218-23.
doi:10.1055/s-2007-970422.
 38. Pi-Sunyer FX, Schweizer A, Mills D, Dejager S. Efficacy and tolerability of vildagliptin monotherapy in drug-naive patients with type 2 diabetes. *Diabetes Res Clin Pract*. 2007;76(1):132-8. doi:10.1016/j.diabres.2006.12.009.
 39. Bosi E, Camisasca RP, Collober C, Rochotte E, Garber AJ. Effects of vildagliptin on glucose control over 24 weeks in patients with type 2 diabetes inadequately controlled with metformin. *Diabetes Care*. 2007;30(4):890-5.
doi:10.2337/dc06-1732.
 40. Garber AJ, Schweizer A, Baron MA, Rochotte E, Dejager S. Vildagliptin in combination with pioglitazone improves glycaemic control in patients with type 2 diabetes failing thiazolidinedione monotherapy: a randomized, placebo-controlled study. *Diabetes Obes Metab*. 2007;9(2):166-74.
doi:10.1111/j.1463-1326.2006.00684.x.
 41. Garber AJ, Foley JE, Banerji MA, Ebeling P, Gudbjornsdottir S, Camisasca RP et al. Effects of vildagliptin on glucose control in patients with type 2 diabetes inadequately controlled with a sulphonylurea. *Diabetes Obes Metab*. 2008;10(11):1047-56. doi:10.1111/j.1463-1326.2008.00859.x.
 42. Kikuchi M, Abe N, Kato M, Terao S, Mimori N, Tachibana H. Vildagliptin dose-dependently improves glycemic control in Japanese patients with type 2 diabetes mellitus. *Diabetes Res Clin Pract*. 2009;83(2):233-40.
doi:10.1016/j.diabres.2008.10.006.
 43. Kikuchi M, Haneda M, Koya D, Tobe K, Onishi Y, Couturier A et al. Efficacy and

- tolerability of vildagliptin as an add-on to glimepiride in Japanese patients with Type 2 diabetes mellitus. *Diabetes Res Clin Pract.* 2010;89(3):216-23.
doi:10.1016/j.diabres.2010.04.017.
44. Kikuchi M, Iwamoto Y, Inagaki N, Yoshioka T, Mimori N, Ebina H. Clinical evaluation of vildagliptin in patients with type 2 diabetes. *Journal of New Remedies & Clinics* 2010;59(2):121-36. [in Japanese]
 45. Pan C, Xing X, Han P, Zheng S, Ma J, Liu J et al. Efficacy and tolerability of vildagliptin as add-on therapy to metformin in Chinese patients with type 2 diabetes mellitus. *Diabetes Obes Metab.* 2012;14(8):737-44.
doi:10.1111/j.1463-1326.2012.01593.x.
 46. Odawara M, Hamada I, Suzuki M. Efficacy and Safety of Vildagliptin as Add-on to Metformin in Japanese Patients with Type 2 Diabetes Mellitus. *Diabetes Ther.* 2014;5(1):169-81. doi:10.1007/s13300-014-0059-x.
 47. Raz I, Hanefeld M, Xu L, Caria C, Williams-Herman D, Khatami H et al. Efficacy and safety of the dipeptidyl peptidase-4 inhibitor sitagliptin as monotherapy in patients with type 2 diabetes mellitus. *Diabetologia.* 2006;49(11):2564-71.
doi:10.1007/s00125-006-0416-z.
 48. Aschner P, Kipnes MS, Lunceford JK, Sanchez M, Mickel C, Williams-Herman DE et al. Effect of the dipeptidyl peptidase-4 inhibitor sitagliptin as monotherapy on glycemic control in patients with type 2 diabetes. *Diabetes Care.* 2006;29(12):2632-7. doi:10.2337/dc06-0703.
 49. Charbonnel B, Karasik A, Liu J, Wu M, Meininger G, Sitagliptin Study G. Efficacy and safety of the dipeptidyl peptidase-4 inhibitor sitagliptin added to ongoing metformin therapy in patients with type 2 diabetes inadequately controlled with metformin alone. *Diabetes Care.* 2006;29(12):2638-43. doi:10.2337/dc06-0706.
 50. Rosenstock J, Brazg R, Andryuk PJ, Lu K, Stein P. Efficacy and safety of the dipeptidyl peptidase-4 inhibitor sitagliptin added to ongoing pioglitazone therapy in patients with type 2 diabetes: a 24-week, multicenter, randomized, double-blind, placebo-controlled, parallel-group study. *Clin Ther.* 2006;28(10):1556-68.
doi:10.1016/j.clinthera.2006.10.007.
 51. Hanefeld M, Herman GA, Wu M, Mickel C, Sanchez M, Stein PP et al. Once-daily sitagliptin, a dipeptidyl peptidase-4 inhibitor, for the treatment of patients with type 2 diabetes. *Curr Med Res Opin.* 2007;23(6):1329-39.
doi:10.1185/030079907X188152.
 52. Goldstein BJ, Feinglos MN, Lunceford JK, Johnson J, Williams-Herman DE, Sitagliptin 036 Study G. Effect of initial combination therapy with sitagliptin, a

- dipeptidyl peptidase-4 inhibitor, and metformin on glycemic control in patients with type 2 diabetes. *Diabetes Care*. 2007;30(8):1979-87. doi:10.2337/dc07-0627.
53. Hermansen K, Kipnes M, Luo E, Fanurik D, Khatami H, Stein P et al. Efficacy and safety of the dipeptidyl peptidase-4 inhibitor, sitagliptin, in patients with type 2 diabetes mellitus inadequately controlled on glimepiride alone or on glimepiride and metformin. *Diabetes Obes Metab*. 2007;9(5):733-45. doi:10.1111/j.1463-1326.2007.00744.x.
 54. Scott R, Wu M, Sanchez M, Stein P. Efficacy and tolerability of the dipeptidyl peptidase-4 inhibitor sitagliptin as monotherapy over 12 weeks in patients with type 2 diabetes. *Int J Clin Pract*. 2007;61(1):171-80. doi:10.1111/j.1742-1241.2006.01246.x.
 55. Raz I, Chen Y, Wu M, Hussain S, Kaufman KD, Amatruda JM et al. Efficacy and safety of sitagliptin added to ongoing metformin therapy in patients with type 2 diabetes. *Curr Med Res Opin*. 2008;24(2):537-50. doi:10.1185/030079908X260925.
 56. Scott R, Loeys T, Davies MJ, Engel SS, Sitagliptin Study G. Efficacy and safety of sitagliptin when added to ongoing metformin therapy in patients with type 2 diabetes. *Diabetes Obes Metab*. 2008;10(10):959-69. doi:10.1111/j.1463-1326.2007.00839.x.
 57. Nonaka K, Kakikawa T, Sato A, Okuyama K, Fujimoto G, Kato N et al. Efficacy and safety of sitagliptin monotherapy in Japanese patients with type 2 diabetes. *Diabetes Res Clin Pract*. 2008;79(2):291-8. doi:10.1016/j.diabres.2007.08.021.
 58. Mohan V, Yang W, Son HY, Xu L, Noble L, Langdon RB et al. Efficacy and safety of sitagliptin in the treatment of patients with type 2 diabetes in China, India, and Korea. *Diabetes Res Clin Pract*. 2009;83(1):106-16. doi:10.1016/j.diabres.2008.10.009.
 59. Iwamoto Y, Taniguchi T, Nonaka K, Okamoto T, Okuyama K, Arjona Ferreira JC et al. Dose-ranging efficacy of sitagliptin, a dipeptidyl peptidase-4 inhibitor, in Japanese patients with type 2 diabetes mellitus. *Endocr J*. 2010;57(5):383-94.
 60. Kashiwagi A, Kadowaki T, Tajima N, Nonaka K, Taniguchi T, Nishii M et al. Sitagliptin added to treatment with ongoing pioglitazone for up to 52 weeks improves glycemic control in Japanese patients with type 2 diabetes. *J Diabetes Investig*. 2011;2(5):381-90. doi:10.1111/j.2040-1124.2011.00120.x.
 61. Barzilai N, Guo H, Mahoney EM, Caporossi S, Golm GT, Langdon RB et al. Efficacy and tolerability of sitagliptin monotherapy in elderly patients with type 2 diabetes: a randomized, double-blind, placebo-controlled trial. *Curr Med Res Opin*. 2011;27(5):1049-58. doi:10.1185/03007995.2011.568059.

62. Tajima N, Kadowaki T, Odawara M, Nishii M, Taniguchi T, Arjona Ferreira JC. Addition of sitagliptin to ongoing glimepiride therapy in Japanese patients with type 2 diabetes over 52 weeks leads to improved glycemic control. *Diabetology International*. 2011;2(1):32-44. doi:10.1007/s13340-011-0022-2.
63. Yang W, Guan Y, Shentu Y, Li Z, Johnson-Levonas AO, Engel SS et al. The addition of sitagliptin to ongoing metformin therapy significantly improves glycemic control in Chinese patients with type 2 diabetes. *J Diabetes*. 2012;4(3):227-37. doi:10.1111/j.1753-0407.2012.00213.x.
64. Nicolle LE, Capuano G, Ways K, Usiskin K. Effect of canagliflozin, a sodium glucose co-transporter 2 (SGLT2) inhibitor, on bacteriuria and urinary tract infection in subjects with type 2 diabetes enrolled in a 12-week, phase 2 study. *Curr Med Res Opin*. 2012;28(7):1167-71. doi:10.1185/03007995.2012.689956.
65. Fonseca V, Staels B, Morgan JD, 2nd, Shentu Y, Golm GT, Johnson-Levonas AO et al. Efficacy and safety of sitagliptin added to ongoing metformin and pioglitazone combination therapy in a randomized, placebo-controlled, 26-week trial in patients with type 2 diabetes. *J Diabetes Complications*. 2013;27(2):177-83. doi:10.1016/j.jdiacomp.2012.09.007.
66. Dobs AS, Goldstein BJ, Aschner P, Horton ES, Umpierrez GE, Duran L et al. Efficacy and safety of sitagliptin added to ongoing metformin and rosiglitazone combination therapy in a randomized placebo-controlled 54-week trial in patients with type 2 diabetes. *J Diabetes*. 2013;5(1):68-79. doi:10.1111/j.1753-0407.2012.00223.x.
67. Roden M, Weng J, Eilbracht J, Delafont B, Kim G, Woerle HJ et al. Empagliflozin monotherapy with sitagliptin as an active comparator in patients with type 2 diabetes: a randomised, double-blind, placebo-controlled, phase 3 trial. *The Lancet Diabetes & Endocrinology*. 2013;1(3):208-19. doi:10.1016/s2213-8587(13)70084-6.
68. Lavalley-Gonzalez FJ, Januszewicz A, Davidson J, Tong C, Qiu R, Canovatchel W et al. Efficacy and safety of canagliflozin compared with placebo and sitagliptin in patients with type 2 diabetes on background metformin monotherapy: a randomised trial. *Diabetologia*. 2013;56(12):2582-92. doi:10.1007/s00125-013-3039-1.
69. Tajima N, Kadowaki T, Okamoto T, Sato A, Okuyama K, Minamide T et al. Sitagliptin added to voglibose monotherapy improves glycemic control in patients with type 2 diabetes. *J Diabetes Investig*. 2013;4(6):595-604. doi:10.1111/jdi.12116.
70. Kadowaki T, Tajima N, Odawara M, Nishii M, Taniguchi T, Ferreira JC. Addition of sitagliptin to ongoing metformin monotherapy improves glycemic control in Japanese patients with type 2 diabetes over 52 weeks. *J Diabetes Investig*.

- 2013;4(2):174-81. doi:10.1111/jdi.12001.
71. Amin NB, Aggarwal N, Pall D, Paragh G, Denney WS, Le V et al. Two dose-ranging studies with PF-04937319, a systemic partial activator of glucokinase, as add-on therapy to metformin in adults with type 2 diabetes. *Diabetes, Obesity and Metabolism*. 2015;17(8):751-9. doi:10.1111/dom.12474.
 72. Ji L, Han P, Wang X, Liu J, Zheng S, Jou YM et al. Randomized clinical trial of the safety and efficacy of sitagliptin and metformin co-administered to Chinese patients with type 2 diabetes mellitus. *J Diabetes Investig*. 2016;7(5):727-36. doi:10.1111/jdi.12511.
 73. Moses RG, Round E, Shentu Y, Golm GT, O'Neill E A, Gantz I et al. A randomized clinical trial evaluating the safety and efficacy of sitagliptin added to the combination of sulfonylurea and metformin in patients with type 2 diabetes mellitus and inadequate glycemic control. *J Diabetes*. 2016;8(5):701-11. doi:10.1111/1753-0407.12351.
 74. Clinical Trial to Evaluate the Safety and Efficacy of the Addition of Sitagliptin in Participants With Type 2 Diabetes Mellitus Receiving Acarbose Monotherapy (MK-0431-130).
<https://clinicaltrials.gov/ct2/show/NCT01177384?term=NCT01177384&rank=1>.
Accessed 7 July 2016
 75. Safety and Efficacy of Empagliflozin (BI 10773) and Sitagliptin Versus Placebo Over 76 Weeks in Patients With Type 2 Diabetes.
<https://clinicaltrials.gov/ct2/show/NCT01289990?term=NCT01289990&rank=1>.
Accessed 7 July 2016
 76. Study of Safety and Efficacy of PF-04991532 in Subjects With Type 2 Diabetes.
<https://clinicaltrials.gov/ct2/show/NCT01338870?term=NCT01338870&rank=1>.
Accessed 7 July 2016
 77. Study Of Safety And Efficacy Of PF-04991532 In Subjects With Type 2 Diabetes Mellitus.
<https://clinicaltrials.gov/ct2/show/NCT01336738?term=NCT01336738&rank=1>.
Accessed 7 July 2016
 78. A Study in China Evaluating the Safety and Efficacy of Adding Sitagliptin to Stable Therapy With Sulfonylurea With or Without Metformin in Participants With Type 2 Diabetes Mellitus (T2DM) (MK-0431-253).
<https://clinicaltrials.gov/ct2/show/NCT01590771?term=NCT01590771&rank=1>.
Accessed 7 July 2016
 79. Omarigliptin (MK-3102) Clinical Trial - Placebo- and Sitagliptin-Controlled

Monotherapy Study in Japanese Patients With Type 2 Diabetes Mellitus
(MK-3102-020).

<https://clinicaltrials.gov/ct2/show/NCT01703221?term=NCT01703221&rank=1>.

Accessed 7 July 2016

Appendix 2

1. List JF, Woo V, Morales E, Tang W, Fiedorek FT. Sodium-glucose cotransport inhibition with dapagliflozin in type 2 diabetes. *Diabetes Care*. 2009;32(4):650-7. doi:10.2337/dc08-1863.
2. Bailey CJ, Gross JL, Pieters A, Bastien A, List JF. Effect of dapagliflozin in patients with type 2 diabetes who have inadequate glycaemic control with metformin: a randomised, double-blind, placebo-controlled trial. *Lancet (London, England)*. 2010;375(9733):2223-33.
3. Ferrannini E, Ramos SJ, Salsali A, Tang W, List JF. Dapagliflozin monotherapy in type 2 diabetic patients with inadequate glycemic control by diet and exercise: a randomized, double-blind, placebo-controlled, phase 3 trial. *Diabetes Care*. 2010;33(10):2217-24. doi:10.2337/dc10-0612.
4. Strojek K, Yoon KH, Hrubá V, Elze M, Langkilde AM, Parikh S. Effect of dapagliflozin in patients with type 2 diabetes who have inadequate glycaemic control with glimepiride: a randomized, 24-week, double-blind, placebo-controlled trial. *Diabetes, obesity & metabolism*. 2011;13(10):928-38. doi:10.1111/j.1463-1326.2011.01434.x.
5. Bailey CJ, Iqbal N, T'Joën C, List JF. Dapagliflozin monotherapy in drug-naive patients with diabetes: a randomized-controlled trial of low-dose range. *Diabetes, obesity & metabolism*. 2012;14(10):951-9. doi:10.1111/j.1463-1326.2012.01659.x.
6. Rosenstock J, Vico M, Wei L, Salsali A, List JF. Effects of dapagliflozin, an SGLT2 inhibitor, on HbA(1c), body weight, and hypoglycemia risk in patients with type 2 diabetes inadequately controlled on pioglitazone monotherapy. *Diabetes Care*. 2012;35(7):1473-8. doi:10.2337/dc11-1693.
7. Kaku K, Inoue S, Matsuoka O, Kiyosue A, Azuma H, Hayashi N et al. Efficacy and safety of dapagliflozin as a monotherapy for type 2 diabetes mellitus in Japanese patients with inadequate glycaemic control: a phase II multicentre, randomized, double-blind, placebo-controlled trial. *Diabetes, obesity & metabolism*. 2013;15(5):432-40. doi:10.1111/dom.12047.
8. Jabbour SA, Hardy E, Sugg J, Parikh S. Dapagliflozin is effective as add-on therapy to sitagliptin with or without metformin: a 24-week, multicenter, randomized, double-blind, placebo-controlled study. *Diabetes Care*. 2014;37(3):740-50. doi:10.2337/dc13-0467.
9. Kaku K, Kiyosue A, Inoue S, Ueda N, Tokudome T, Yang J et al. Efficacy and safety of dapagliflozin monotherapy in Japanese patients with type 2 diabetes

- inadequately controlled by diet and exercise. *Diabetes, obesity & metabolism*. 2014;16(11):1102-10. doi:10.1111/dom.12325.
10. Ji L, Ma J, Li H, Mansfield TA, T'Joen C L, Iqbal N et al. Dapagliflozin as monotherapy in drug-naive Asian patients with type 2 diabetes mellitus: a randomized, blinded, prospective phase III study. *Clinical therapeutics*. 2014;36(1):84-100 e9. doi:10.1016/j.clinthera.2013.11.002.
 11. Matthaeci S, Bowering K, Rohwedder K, Grohl A, Parikh S. Dapagliflozin Improves Glycemic Control and Reduces Body Weight as Add-on Therapy to Metformin Plus Sulfonylurea: A 24-Week Randomized, Double-Blind Clinical Trial. *Diabetes Care*. 2015;38(3):365-72. doi:10.2337/dc14-0666.
 12. Schumm-Draeger PM, Burgess L, Koranyi L, Hruba V, Hamer-Maansson JE, de Bruin TW. Twice-daily dapagliflozin co-administered with metformin in type 2 diabetes: a 16-week randomized, placebo-controlled clinical trial. *Diabetes, obesity & metabolism*. 2015;17(1):42-51. doi:10.1111/dom.12387.
 13. Safety and Efficacy of Dapagliflozin as Monotherapy in Subjects With Type 2 Diabetes.
<https://www.clinicaltrials.gov/ct2/show/NCT00736879?term=NCT00736879&rank=1>. Accessed 14 May 2017
 14. Rosenstock J, Aggarwal N, Polidori D, Zhao Y, Arbit D, Usiskin K et al. Dose-ranging effects of canagliflozin, a sodium-glucose cotransporter 2 inhibitor, as add-on to metformin in subjects with type 2 diabetes. *Diabetes Care*. 2012;35(6):1232-8. doi:10.2337/dc11-1926.
 15. Wilding JP, Charpentier G, Hollander P, Gonzalez-Galvez G, Mathieu C, Vercruysse F et al. Efficacy and safety of canagliflozin in patients with type 2 diabetes mellitus inadequately controlled with metformin and sulphonylurea: a randomised trial. *International journal of clinical practice*. 2013;67(12):1267-82. doi:10.1111/ijcp.12322.
 16. Inagaki N, Kondo K, Yoshinari T, Maruyama N, Susuta Y, Kuki H. Efficacy and safety of canagliflozin in Japanese patients with type 2 diabetes: a randomized, double-blind, placebo-controlled, 12-week study. *Diabetes, obesity & metabolism*. 2013;15(12):1136-45. doi:10.1111/dom.12149.
 17. Lavalley-Gonzalez FJ, Januszewicz A, Davidson J, Tong C, Qiu R, Canovatchel W et al. Efficacy and safety of canagliflozin compared with placebo and sitagliptin in patients with type 2 diabetes on background metformin monotherapy: a randomised trial. *Diabetologia*. 2013;56(12):2582-92. doi:10.1007/s00125-013-3039-1.
 18. Stenlof K, Cefalu WT, Kim KA, Alba M, Usiskin K, Tong C et al. Efficacy and

- safety of canagliflozin monotherapy in subjects with type 2 diabetes mellitus inadequately controlled with diet and exercise. *Diabetes, obesity & metabolism*. 2013;15(4):372-82. doi:10.1111/dom.12054.
19. Forst T, Guthrie R, Goldenberg R, Yee J, Vijapurkar U, Meininger G et al. Efficacy and safety of canagliflozin over 52 weeks in patients with type 2 diabetes on background metformin and pioglitazone. *Diabetes, obesity & metabolism*. 2014;16(5):467-77. doi:10.1111/dom.12273.
 20. Inagaki N, Kondo K, Yoshinari T, Takahashi N, Susuta Y, Kuki H. Efficacy and safety of canagliflozin monotherapy in Japanese patients with type 2 diabetes inadequately controlled with diet and exercise: a 24-week, randomized, double-blind, placebo-controlled, Phase III study. *Expert opinion on pharmacotherapy*. 2014;15(11):1501-15. doi:10.1517/14656566.2014.935764.
 21. Qiu R, Capuano G, Meininger G. Efficacy and safety of twice-daily treatment with canagliflozin, a sodium glucose co-transporter 2 inhibitor, added on to metformin monotherapy in patients with type 2 diabetes mellitus. *Journal of Clinical & Translational Endocrinology*. 2014;1(2):54-60. doi:10.1016/j.jcte.2014.04.001.
 22. Ji L, Han P, Liu Y, Yang G, Dieu Van NK, Vijapurkar U et al. Canagliflozin in Asian patients with type 2 diabetes on metformin alone or metformin in combination with sulphonylurea. *Diabetes, obesity & metabolism*. 2015;17(1):23-31. doi:10.1111/dom.12385.
 23. Haring HU, Merker L, Seewaldt-Becker E, Weimer M, Meinicke T, Woerle HJ et al. Empagliflozin as add-on to metformin plus sulfonylurea in patients with type 2 diabetes: a 24-week, randomized, double-blind, placebo-controlled trial. *Diabetes Care*. 2013;36(11):3396-404. doi:10.2337/dc12-2673.
 24. Roden M, Weng J, Eilbracht J, Delafont B, Kim G, Woerle HJ et al. Empagliflozin monotherapy with sitagliptin as an active comparator in patients with type 2 diabetes: a randomised, double-blind, placebo-controlled, phase 3 trial. *The Lancet Diabetes & Endocrinology*. 2013;1(3):208-19. doi:10.1016/s2213-8587(13)70084-6.
 25. Rosenstock J, Seman LJ, Jelaska A, Hantel S, Pinnetti S, Hach T et al. Efficacy and safety of empagliflozin, a sodium glucose cotransporter 2 (SGLT2) inhibitor, as add-on to metformin in type 2 diabetes with mild hyperglycaemia. *Diabetes, obesity & metabolism*. 2013;15(12):1154-60. doi:10.1111/dom.12185.
 26. Ferrannini E, Seman L, Seewaldt-Becker E, Hantel S, Pinnetti S, Woerle HJ. A Phase IIB, randomized, placebo-controlled study of the SGLT2 inhibitor empagliflozin in patients with type 2 diabetes. *Diabetes, obesity & metabolism*.

- 2013;15(8):721-8. doi:10.1111/dom.12081.
27. Kovacs CS, Seshiah V, Swallow R, Jones R, Rattunde H, Woerle HJ et al. Empagliflozin improves glycaemic and weight control as add-on therapy to pioglitazone or pioglitazone plus metformin in patients with type 2 diabetes: a 24-week, randomized, placebo-controlled trial. *Diabetes, obesity & metabolism*. 2014;16(2):147-58. doi:10.1111/dom.12188.
 28. Haring HU, Merker L, Seewaldt-Becker E, Weimer M, Meinicke T, Broedl UC et al. Empagliflozin as add-on to metformin in patients with type 2 diabetes: a 24-week, randomized, double-blind, placebo-controlled trial. *Diabetes Care*. 2014;37(6):1650-9. doi:10.2337/dc13-2105.
 29. Kadowaki T, Haneda M, Inagaki N, Terauchi Y, Taniguchi A, Koiwai K et al. Empagliflozin monotherapy in Japanese patients with type 2 diabetes mellitus: a randomized, 12-week, double-blind, placebo-controlled, phase II trial. *Advances in therapy*. 2014;31(6):621-38. doi:10.1007/s12325-014-0126-8.
 30. Ross S, Thamer C, Cescutti J, Meinicke T, Woerle HJ, Broedl UC. Efficacy and safety of empagliflozin twice daily versus once daily in patients with type 2 diabetes inadequately controlled on metformin: a 16-week, randomized, placebo-controlled trial. *Diabetes, obesity & metabolism*. 2015;17(7):699-702. doi:10.1111/dom.12469.
 31. Roden M, Merker L, Christiansen AV, Roux F, Salsali A, Kim G et al. Safety, tolerability and effects on cardiometabolic risk factors of empagliflozin monotherapy in drug-naive patients with type 2 diabetes: a double-blind extension of a Phase III randomized controlled trial. *Cardiovasc Diabetol*. 2015;14(1):154. doi:10.1186/s12933-015-0314-0.
 32. Haering HU, Merker L, Christiansen AV, Roux F, Salsali A, Kim G et al. Empagliflozin as add-on to metformin plus sulphonylurea in patients with type 2 diabetes. *Diabetes research and clinical practice*. 2015;110(1):82-90. doi:10.1016/j.diabres.2015.05.044.
 33. Safety and Efficacy of Empagliflozin (BI 10773) and Sitagliptin Versus Placebo Over 76 Weeks in Patients With Type 2 Diabetes. <https://www.clinicaltrials.gov/ct2/show/NCT01289990?term=NCT01289990&rank=1>. Accessed 14 May 2017